A note on sharp identification regions



Let $P := \{ \mathbb{P}_{\theta} \mid \theta \in \Theta \}$ be a statistical model and

- Y,... unobservable random variables,
- X, Y, Y, ... observable random variables w.r.t an underlying probability space (Ω, F, P).
- The joint distribution of the random Variables X, Y, Y, Y under a model P_θ is denoted with F_θ and the joint distribution under the "true model" P is denoted with F^{X,Y,Y,Y}.
- The unobserved variables fullfill a certain condition C(X, Y, Y, Y) = 1.
 e.g. Y ≤ Y ≤ Y or ∀X : E(Y | X) ≤ E(Y | X) ≤ E(Y | X).

• Two parameters θ_1 and θ_2 are undistinguishable (i.e. $\theta_1 \sim \theta_2$) if the corresponding models \mathbb{P}_{θ_1} and \mathbb{P}_{θ_2} are empirically undistinguishable, which means, that the distributions of the observable variables are the same:

$$\mathcal{F}_{\theta_{1}}^{X,\underline{Y},\overline{Y}} = \mathcal{F}_{\theta_{2}}^{X,\underline{Y},\overline{Y}}$$

A statistical model P is called point-identified, if any two different parameters θ_1 and θ_2 are empirically distinguishable, i.e.:

$$\sim = \Delta_{\Theta} = \{(\theta, \theta) \mid \theta \in \Theta\}.$$

Otherwise it is called partially identified.

Example

The simple linear model

$$\Theta = B imes \mathbb{R}_{\geq 0} imes \mathcal{Z}(\mathbb{R}_{\geq 0}) imes \mathcal{Z}(\mathbb{R}_{\geq 0})$$

with $B = \mathbb{R}^2$. For $\theta = (\beta, \sigma^2, \sigma_I, \sigma_u) \in \Theta$, the random variables are defined as:

$$Y = X\beta + \varepsilon$$

$$Y = X\beta + \varepsilon - \sigma_{l}$$

$$\overline{Y} = X\beta + \varepsilon + \sigma_{u}$$

with $\varepsilon \sim N(0, \sigma^2 I)$.

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with $\varepsilon \sim N(0, \sigma^2 I)$.

Here we are only interested in the values of $\beta \in B$.

This model is only partially identified. For example

 $((\beta_0, \beta_1), \sigma^2, 0, 1) \sim ((\beta_0 + 1, \beta_1), \sigma^2, 1, 0).$



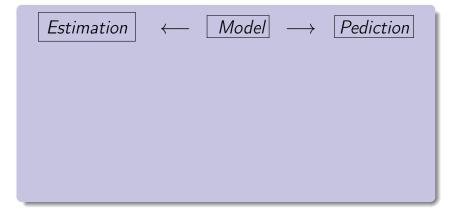
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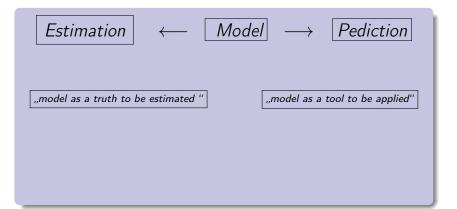
Moreover, the quotient space $\Theta_{/\sim}$ ist not of the form

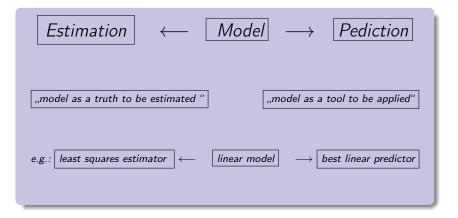
$$\Theta_{/\sim} = B_{/\approx} \times , rest^{\prime\prime},$$

so we must factorize the whole space Θ and not only the interesting B to make the model point-identified.



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$$\begin{array}{l} Y \sim F_{\theta} \\ \Longleftrightarrow \quad F^{Y} = F_{\theta}^{Y} \\ \Leftrightarrow \quad L(F_{\theta}^{Y}, F) = 0 \end{array}$$

for some distance-function $L(\cdot, \cdot)$.

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The actual problem is, that F^{Y} is unknown \Longrightarrow later.

Let $P = \{\mathbb{P}_{\theta} \mid \theta \in \Theta\}$ be a statistical model with the corresponding joint distributions $\{F_{\theta}^{X,Y,\underline{Y},\overline{Y}} \mid \theta \in \Theta\}$ and $X, \underline{Y}, \overline{Y}$ random variables with the joint distribution $F^{X,\underline{Y},\overline{Y}}$. The **Sharp Estimation Region (SER)** is defined as:

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$$SER(\underline{Y}, \overline{Y}) = \underset{\theta \in \Theta}{\operatorname{argmin}} \left(\inf_{\substack{Y \in .t. C(X, Y, \underline{Y}, \overline{Y}) = 1}} L\left(F_{\theta}, F^{X, Y, \overline{Y}, \underline{Y}}\right) \right)$$

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$$SPR(\underline{Y}, \overline{Y}) := \left\{ \operatorname{argmin}_{\theta \in \Theta} L\left(F_{\theta}, F^{X, Y, \overline{Y}, \underline{Y}}\right) \mid Y \text{ s.t. } C(X, Y, \underline{Y}, \overline{Y}) = 1 \right\}$$

We are only interested in the components (β_0, β_1) of an element $\theta = ((\beta_0, \beta_1), \sigma^2, \sigma_l, \sigma_u) \in SER$ and denote the set

$$\{(\beta_0,\beta_1) \mid ((\beta_0,\beta_1),\sigma^2,\sigma_l,\sigma_u) \in SER\}$$

as the sharp estimation region (analogously for the sharp prediction region).

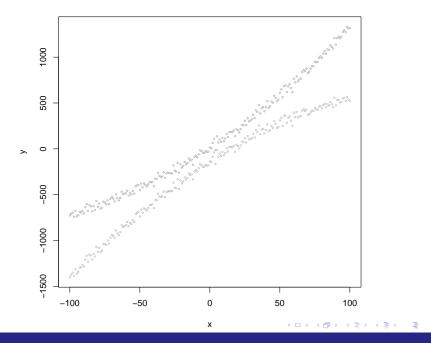
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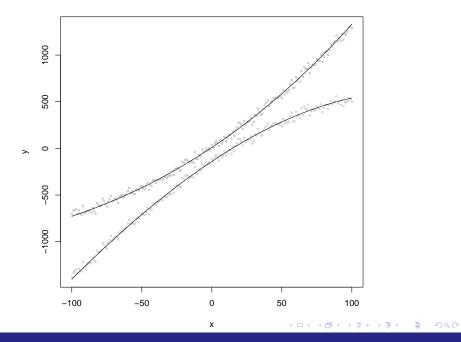
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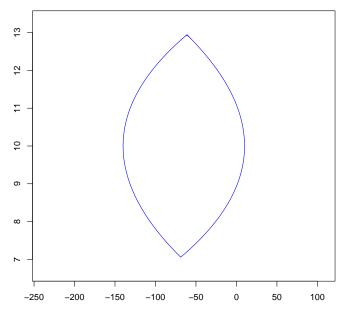
A note on sharp identification regions

$$SER = \{\beta \in B | \mathbb{E}(\underline{Y} | X) \le X\beta \le \mathbb{E}(\overline{Y} | X) \}$$
$$SPR = \{ \operatorname*{argmin}_{\beta \in B} \mathbb{E}((X\beta - Y)^2) | Y \in [\underline{Y}, \overline{Y}] \}$$
$$= \{ (X'X)^{-1}X'Y | Y \in [\underline{Y}, \overline{Y}] \}$$

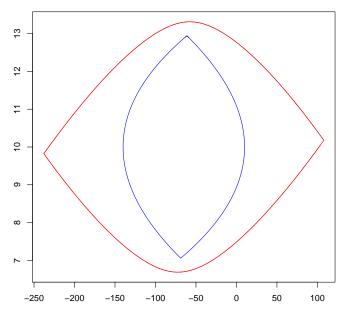
Linear Model



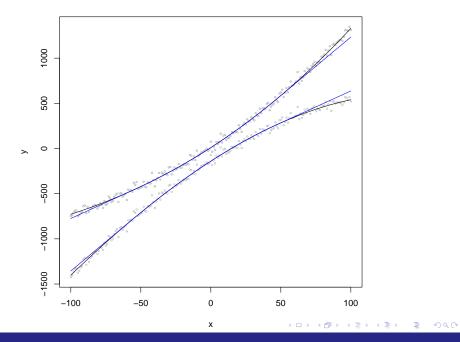


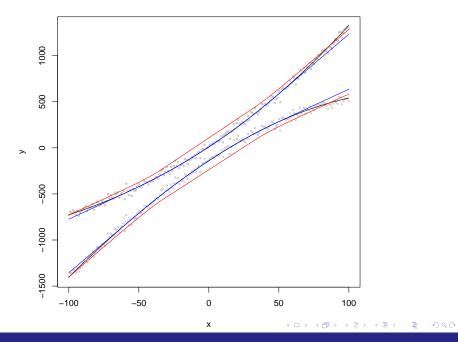


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Theorem

Let $I \subset \mathbb{R}^2$ be a compact convex set. Then there exist random variables $X, \underline{Y}, \overline{Y}$ such that

 $SER(X, \underline{Y}, \overline{Y}) = I$,

namely:

$$X \sim N(0,1)$$

$$\underline{Y} = \min\{\beta_0 + \beta_1 X \mid (\beta_0, \beta_1) \in I\}$$

$$\underline{Y} = \max\{\beta_0 + \beta_1 X \mid (\beta_0, \beta_1) \in I\}$$

The Minkowski-Sum

$$M = \bigoplus_{i=1}^n l_i = \left\{ \sum_{i=1}^n p_i \mid p_i \in l_i \right\}$$

of *n* line-segments $l_i \subseteq \mathbb{R}^d$ is called a **zonotope**.

A zonotope is a convex, compact and centrally symmetric polytope with finite many extremepoints and central-symmetric facets.

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Definition

A closed, centrally symmetric convex set $Z \subseteq \mathbb{R}^d$ is called a **zonoid**, if it can be approximated arbitrarily closely by zonotopes (w.r.t. a metric, e.g. the Hausdorff distance). For d = 2 the zonoids are exactly the closed, centrally symmetric convex sets.

Let $I \subseteq \mathbb{R}^2$ be a zonoid in general position. Then there exists random variables $X, \underline{Y}, \overline{Y}$ such that

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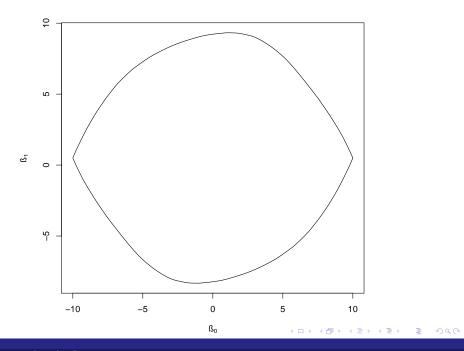
Lemma

Let $I = SPR(X, \underline{Y}^*, \overline{Y}^*) \subseteq \mathbb{R}^2$ be a zonoid and $E \subseteq SER(X, \underline{Y}^*, \overline{Y}^*)$ an arbitrary compact convex set. Then for every $\varepsilon > 0$ there exist random variables $X, \underline{Y}, \overline{Y}$ such that:

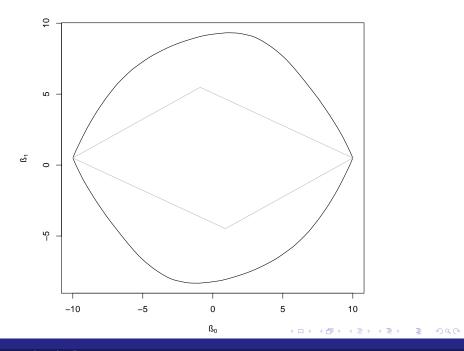
 $d_{H}(SPR(X,\underline{Y},\overline{Y}),I) \leq \varepsilon$ $d_{H}(SER(X,\underline{Y},\overline{Y}),E) \leq \varepsilon$

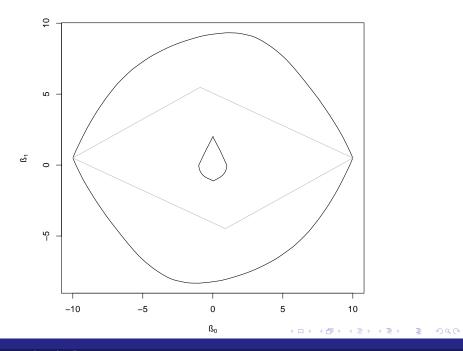
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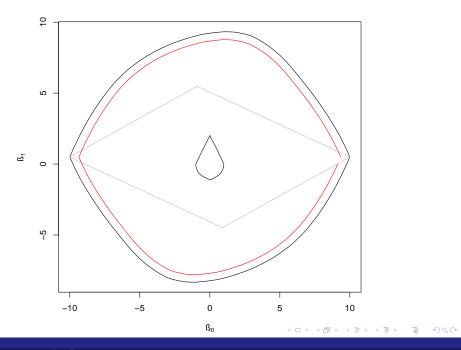
with the Hausdorff distance d_H .

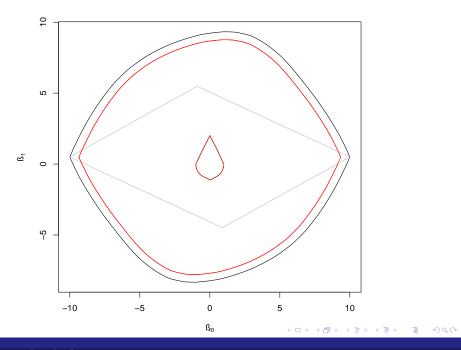


A note on sharp identification regions









Mappings between ordered sets

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Mappings between ordered sets

Definition

Let (P, \leq) and (Q, \sqsubseteq) be partially ordered sets. A pair (f, g) of mappings $f : P \longrightarrow Q$ and $g : Q \longrightarrow P$ is called **adjunction**, if:

$$\forall p \in P \forall q \in Q: p \leq g(q) \iff f(p) \sqsubseteq q.$$

In this case, f is called **left adjoint** and g is called **right adjoint**.

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• Dempster-Shafer-Theory: Multivalued mapping $\Gamma : X \longrightarrow 2^S$ with corresponding

$$\widetilde{\Gamma}:(2^X,\subseteq)\longrightarrow(2^S,\subseteq):A\mapsto \bigcup_{a\in A}\Gamma(a)$$
 and the operator

 $_*: (2^S, \subseteq) \longrightarrow (2^X, \subseteq): T \mapsto \{x \in X \mid \Gamma(x) \subseteq T\}.$

The pair $(\tilde{\Gamma}, *)$ is an adjunction. From this, the ∞ -monotonicity of a Belief-function

$$Bel = P \circ *$$

with P a probability-measure follows immediately, since P is ∞ -monotone and * is meet-preserving. Furthermure it is clear, that also Bel $\circ *$ is ∞ -monotone.

• Lower coherent previsions:

$$f : \underline{P} \mapsto \mathcal{M}(\underline{P}) = \{ p \in \mathscr{P}(\Omega) \mid p \ge \underline{P} \}$$
 and
 $g : M \mapsto \underline{P}_M : X \mapsto \inf_{p \in M} p(X)$ are an adjunction.

 Formal concept analysis: Incidence structure $\mathbb{K} = (G, M, I)$ with $G \ldots$ objects, $M \ldots$ attributes and a relation $I \subseteq G \times M$. $(g, m) \in I$ means object g has attribute m (also denoted as glm). $f: (2^M, \subset) \longrightarrow (2^G, \subset): X \mapsto \{g \in G | \forall m \in X : glm\}$ "The set of all objects having all attributes in X" $g: (2^G, \supseteq) \longrightarrow (2^M, \supseteq): Y \mapsto \{m \in M \mid \forall g \in Y : glm\}$ "...The set of all joint attributes of all objects in Y". The pair (f, g) is an adjunction.

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- If P is a complete lattice, than f is a left adjoint, if and only if f is join-preserving.
- If Q is a complete lattice, than g is a right adjoint, if and only if g is meet-preserving.

The mapping

$$\begin{array}{ll} \textit{SER} & : & (\mathcal{Z}(\Omega), \leq) \longrightarrow (2^{\mathcal{B}}, \subseteq) : (X, \underline{Y}, \overline{Y}) \mapsto \{\beta \mid \mathbb{E}(\underline{Y} \mid X) \leq \beta X \leq \mathbb{E}(\overline{Y} \mid x)\} \\ \textit{with} \end{array}$$

$$\begin{array}{|c|c|c|c|c|}\hline (X_1,\underline{Y}_1,\overline{Y}_1) \leq (X_2,\underline{Y}_2,\overline{Y}_2) \vdots & \iff & \hline \mathbb{E}(\underline{Y}_1 \mid X) \geq \mathbb{E}(\underline{Y}_2 \mid X) \And \mathbb{E}(\overline{Y}_1 \mid X) \leq \mathbb{E}(\overline{Y}_2 \mid X) \\ & i.e.: (X_1,\underline{Y}_1,\overline{Y}_1) \text{ is more precise} \\ & & than (X_2,\underline{Y}_2,\overline{Y}_2) \end{array}$$

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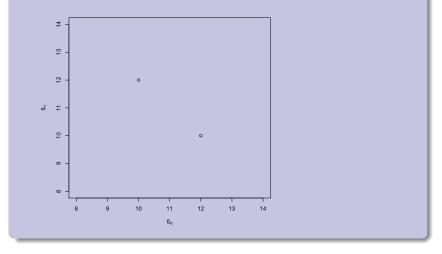
is a right adjoint. The corresponding left adjoint is the "prediction-operator":

$$PR: \quad (2^{\mathcal{B}}, \subseteq) \longrightarrow \left(\mathcal{Z}(\Omega), \leq) : M \mapsto (X, \min_{\beta \in M} X\beta, \max_{\beta \in M} X\beta) \right)$$

Thus, the following holds:

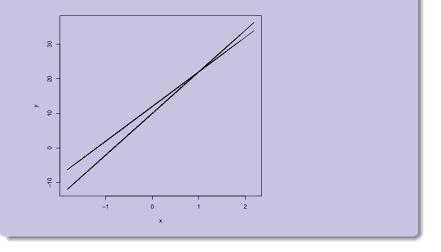
- A1 SER o PR is extensive and PR o SER is intensive.
- A2 PR and SER are order-preserving.
- A3 $PR \circ SER \circ PR = PR$ and $SER \circ PR \circ SER = SER$ and thus $PR \circ SER$ and $SER \circ PR$ are idempotent.
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- A5 The adjoints PR and SER are determining each other unambiguously.
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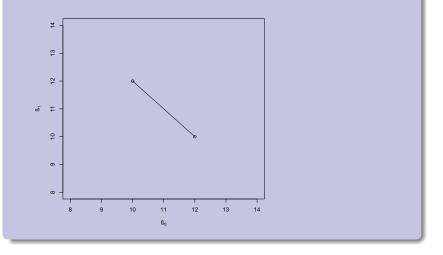


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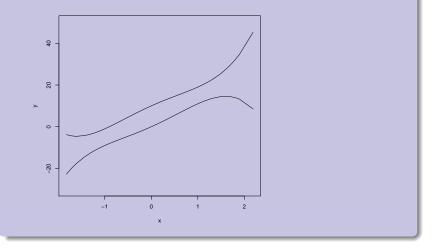


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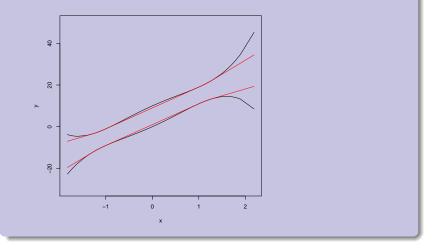
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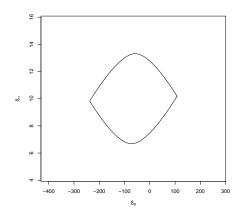
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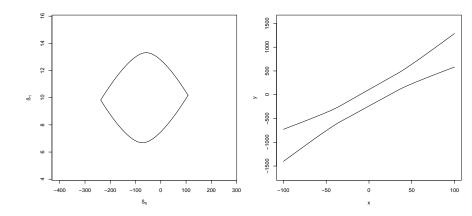
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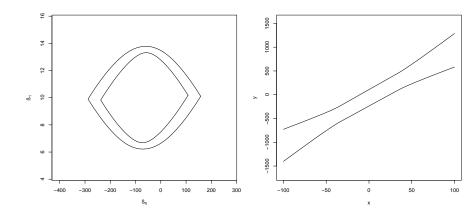
The mapping $SPR : (\mathcal{Z}(\Omega), \leq) \longrightarrow (2^{\mathcal{B}}, \subseteq) : (X, \underline{Y}, \overline{Y}) \mapsto \{(X'X)^{-1}X'Y \mid \underline{Y} \leq Y \leq \overline{Y}\}$ is no right adjoint, since it is not meet-preserving. In general $SPR(Z_1 \land Z_2) \neq SPR(Z_1) \cap SPR(Z_2)$, since the intersection of two zonoids is in general not a zonoid. Thus, in general, only $SPR \circ PR \circ SPR \supset SPR$ holds.



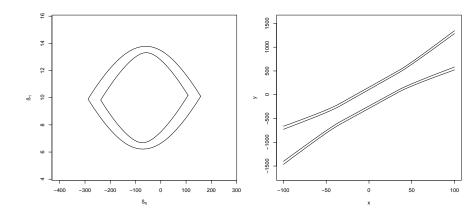
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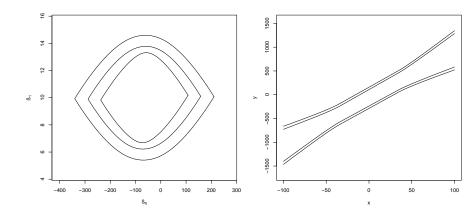
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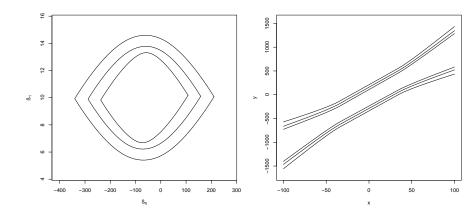


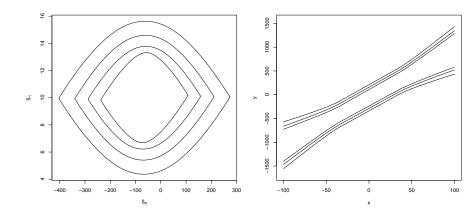
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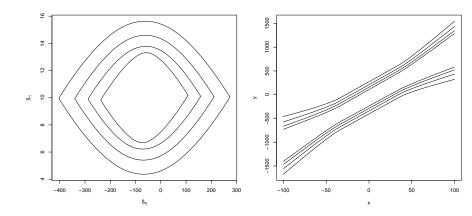
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Definition

Let $E : (P, \leq) \longrightarrow (Q, \sqsubseteq)$ be a mapping. The monotone hull of E is defined as:

$$H(E)$$
 : $(P, \leq) \longrightarrow (Q, \sqsubseteq) : X \mapsto \bigvee_{Y \leq X} E(Y).$

The monotone kernel of E is defined as:

$$K(E)$$
 : $(P, \leq) \longrightarrow (Q, \sqsubseteq) : X \mapsto \bigwedge_{Y \geq X} E(Y).$

These set-valued mappings are both order-preserving (i.e: $X \le Y \Longrightarrow (H(E))(X) \sqsubseteq (H(E))(Y)$ & $(K(E))(X) \sqsubseteq (K(E))(Y)$).

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Lemma

Let the criterion-function $Q: B \longrightarrow \mathbb{R}$ be defined as:

$$Q(\beta) = \int \left\{ \left(\mathbb{E}(\underline{Y} | x) - x\beta \right)_{+}^{2} + \left(\mathbb{E}(\overline{Y} | x) - x\beta \right)_{-}^{2} \right\} d\mathbb{P}(x).$$

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Then the criterion-based mapping

$$E_Q: \mathcal{Z}(\Omega) \longrightarrow 2^B: (X, \underline{Y}, \overline{Y}) \mapsto \operatorname*{argmin}_{\beta \in B} Q(\beta)$$

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$$SPR = H(E_Q)$$

 $SER = K(E_Q).$

Estimation of SER and SPR

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In general, there is no consistent and (in a certain sense) robust estimator of SER.

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