## A note on sharp identification regions

## Definition

Let $P:=\left\{\mathbb{P}_{\theta} \mid \theta \in \Theta\right\}$ be a statsitical model and

- $Y, \ldots$ unobservable random variables,
- X, $\underline{Y}, \bar{Y}, \ldots$ observable random variables w.r.t an underlying probability space $(\Omega, \mathcal{F}, \mathbb{P})$.
- The joint distribution of the random Variables $X, Y, \underline{Y}, \bar{Y}$ under a model $P_{\theta}$ is denoted with $F_{\theta}$ and the joint distribution under the ,,true model" $\mathbb{P}$ is denoted with $F^{X, Y, \underline{Y}, \bar{Y}}$.
- The unobserved variables fullfill a certain condition $C(X, Y, \underline{Y}, \bar{Y})=1$. e.g. $\underline{Y} \leq Y \leq \bar{Y} \quad$ or $\quad \forall X: \mathbb{E}(\underline{Y} \mid X) \leq \mathbb{E}(Y \mid X) \leq \mathbb{E}(\bar{Y} \mid X)$.


## Definition

- Two parameters $\theta_{1}$ and $\theta_{2}$ are undistinguishable (i.e. $\theta_{1} \sim \theta_{2}$ ) if the corresponding models $\mathbb{P}_{\theta_{1}}$ and $\mathbb{P}_{\theta_{2}}$ are empirically undistinguishable, which means, that the distributions of the observable variables are the same:

$$
F_{\theta_{1}}^{X, \underline{Y}, \bar{Y}}=F_{\theta_{2}}^{X, \underline{Y}, \bar{Y}} .
$$

## Definition

A statistical model $P$ is called point-identified, if any two different parameters $\theta_{1}$ and $\theta_{2}$ are empirically distinguishable, i.e.:

$$
\sim=\Delta_{\Theta}=\{(\theta, \theta) \mid \theta \in \Theta\}
$$

Otherwise it is called partially identified.

## Example

The simple linear model

$$
\Theta=B \times \mathbb{R}_{\geq 0} \times \mathcal{Z}\left(\mathbb{R}_{\geq 0}\right) \times \mathcal{Z}\left(\mathbb{R}_{\geq 0}\right)
$$

with $B=\mathbb{R}^{2}$. For $\theta=\left(\beta, \sigma^{2}, \sigma_{l}, \sigma_{u}\right) \in \Theta$, the random variables are defined as:

$$
\begin{aligned}
& Y=X \beta+\varepsilon \\
& \underline{Y}=X \beta+\varepsilon-\sigma_{I} \\
& \overline{\mathbf{Y}}=X \beta+\varepsilon+\sigma_{u}
\end{aligned}
$$

with $\varepsilon \sim N\left(0, \sigma^{2} I\right)$.

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with $\varepsilon \sim N\left(0, \sigma^{2} I\right)$.
Here we are only interested in the values of $\beta \in B$.

This model is only partially identified. For example

$$
\left(\left(\beta_{0}, \beta_{1}\right), \sigma^{2}, 0,1\right) \quad \sim\left(\left(\beta_{0}+1, \beta_{1}\right), \sigma^{2}, 1,0\right)
$$

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Moreover, the quotient space $\Theta_{/ \sim}$ ist not of the form

$$
\Theta_{/ \sim}=B / \approx \times{ }^{\prime} \text { rest", }
$$

so we must factorize the whole space $\Theta$ and not only the interesting $B$ to make the model point-identified.

## Estimation <br>  <br> Model $\longrightarrow$ Pediction

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 ,,model as a truth to be estimated ,,model as a tool to be applied"
# Estimation <br>  <br> Model $\longrightarrow$ Pediction 

,,model as a truth to be estimated
,,model as a tool to be applied"
e.g.: least squares estimator $\longleftarrow \quad$ linear model $\longrightarrow$ best linear predictor

## ＂Estimation＂

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$$
\begin{array}{ll} 
& Y \sim F_{\theta} \\
\Longleftrightarrow & F^{Y}=F_{\theta}^{Y} \\
\Longleftrightarrow & L\left(F_{\theta}^{Y}, F\right)=0
\end{array}
$$

for some distance-function $L(\cdot, \cdot)$.

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- also makes sense, if $F^{Y} \notin\left\{F_{\theta}^{Y} \mid \theta \in \Theta\right\}$.


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Given $F^{Y}$ of the class $\left\{F_{\theta}^{Y} \mid \theta \in \Theta\right\}$,
find (all) $\theta$, such that

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L\left(F_{\theta}, F^{\curlyvee}\right)
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is minimal.

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- if the model is correctly specified, then „prediction" and „estimation" are ,,nearly the same".


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The actual problem is, that $F^{\curlyvee}$ is unknown $\Longrightarrow$ later.

## Definition

Let $P=\left\{\mathbb{P}_{\theta} \mid \theta \in \Theta\right\}$ be a statistical model with the corresponding joint distributions $\left\{F_{\theta}^{X, Y, \underline{Y}, \bar{Y}} \mid \theta \in \Theta\right\}$ and $X, \underline{Y}, \bar{Y}$ random variables with the joint distribution $F^{X, \underline{Y}, \bar{Y}}$. The Sharp Estimation Region (SER) is defined as:

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If the model is correctly specified, this region can also be written as:

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\operatorname{SER}(\underline{Y}, \bar{Y})=\underset{\theta \in \Theta}{\operatorname{argmin}}\left(\inf _{Y_{s . t} . C(X, Y, \underline{Y}, \bar{Y})=1} L\left(F_{\theta}, F^{X, Y, \bar{Y}, \underline{Y}}\right)\right) .
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\operatorname{SPR}(\underline{Y}, \bar{Y}):=\left\{\underset{\theta \in \Theta}{\operatorname{argmin}} L\left(F_{\theta}, F^{X, Y, \bar{Y}, \underline{Y}}\right) \mid Y \text { s.t. } C(X, Y, \underline{Y}, \bar{Y})=1\right\} .
$$

## Now: Linear Model

We are only interested in the components $\left(\beta_{0}, \beta_{1}\right)$ of an element $\theta=\left(\left(\beta_{0}, \beta_{1}\right), \sigma^{2}, \sigma_{l}, \sigma_{u}\right) \in S E R$ and denote the set

$$
\left\{\left(\beta_{0}, \beta_{1}\right) \mid\left(\left(\beta_{0}, \beta_{1}\right), \sigma^{2}, \sigma_{l}, \sigma_{u}\right) \in S E R\right\}
$$

as the sharp estimation region (analogously for the sharp prediction region).

## Linear Model

$$
\begin{aligned}
S E R & =\{\beta \in B \mid \mathbb{E}(\underline{Y} \mid X) \leq X \beta \leq \mathbb{E}(\overline{\mathrm{Y}} \mid X)\} \\
S P R & =\left\{\underset{\beta \in B}{\operatorname{argmin}} \mathbb{E}\left((X \beta-Y)^{2}\right) \mid Y \in[\underline{Y}, \bar{Y}]\right\} \\
& =\left\{\left(X^{\prime} X\right)^{-1} X^{\prime} Y \mid Y \in[\underline{Y}, \bar{Y}]\right\}
\end{aligned}
$$




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Theorem
Let $I \subset \mathbb{R}^{2}$ be a compact convex set．Then there exist random variables $X, \underline{Y}, \bar{Y}$ such that

$$
\operatorname{SER}(X, \underline{Y}, \bar{Y})=I
$$

namely：

$$
\begin{aligned}
& X \sim N(0,1) \\
& \underline{Y}=\min \left\{\beta_{0}+\beta_{1} X \mid\left(\beta_{0}, \beta_{1}\right) \in I\right\} \\
& \underline{Y}=\max \left\{\beta_{0}+\beta_{1} X \mid\left(\beta_{0}, \beta_{1}\right) \in I\right\}
\end{aligned}
$$

## Definition

The Minkowski-Sum

$$
M=\bigoplus_{i=1}^{n} I_{i}=\left\{\sum_{i=1}^{n} p_{i} \mid p_{i} \in I_{i}\right\}
$$

of $n$ line-segments $l_{i} \subseteq \mathbb{R}^{d}$ is called a zonotope.
A zonotope is a convex, compact and centrally symmetric polytope with finite many extremepoints and central-symmetric facets.

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## Definition

A closed, centrally symmetric convex set $Z \subseteq \mathbb{R}^{d}$ is called a zonoid, if it can be approximated arbitrarily closely by zonotopes (w.r.t. a metric, e.g. the Hausdorff distance).
For $d=2$ the zonoids are exactly the closed, centrally symmetric convex sets.

## Lemma

Let $I \subseteq \mathbb{R}^{2}$ be a zonoid in general position. Then there exists random variables $X, \underline{Y}, \bar{Y}$ such that

$$
\operatorname{SPR}(X, \underline{Y}, \bar{Y})=I
$$

## Lemma

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$$
\operatorname{SPR}(X, \underline{Y}, \bar{Y})=1 .
$$

## Lemma

Let $I=\operatorname{SPR}\left(X, \underline{Y}^{*}, \bar{Y}^{*}\right) \subseteq \mathbb{R}^{2}$ be a zonoid and $E \subseteq \operatorname{SER}\left(X, \underline{Y}^{*}, \overline{\mathrm{Y}}^{*}\right)$ an arbitrary compact convex set．Then for every $\varepsilon>0$ there exist random variables $X, \underline{Y}, \bar{Y}$ such that：

$$
\begin{aligned}
d_{H}(\operatorname{SPR}(X, \underline{Y}, \bar{Y}), I) & \leq \varepsilon \\
d_{H}(\operatorname{SER}(X, \underline{Y}, \bar{Y}), E) & \leq \varepsilon
\end{aligned}
$$

with the Hausdorff distance $d_{H}$ ．


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## Mappings between ordered sets

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## Definition

Let $(P, \leq)$ and $(Q, \sqsubseteq)$ be partially ordered sets. A pair $(f, g)$ of mappings $f: P \longrightarrow Q$ and $g: Q \longrightarrow P$ is called adjunction, if:

$$
\forall p \in P \forall q \in Q: \quad p \leq g(q) \Longleftrightarrow f(p) \sqsubseteq q .
$$

In this case, $f$ is called left adjoint and $g$ is called right adjoint.

## Examples of adjunctions

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- Dempster-Shafer-Theory:

Multivalued mapping $\Gamma: X \longrightarrow 2^{S}$ with corresponding
$\tilde{\Gamma}:\left(2^{X}, \subseteq\right) \longrightarrow\left(2^{S}, \subseteq\right): A \mapsto \bigcup_{a \in A} \Gamma(a)$ and the operator
$*:\left(2^{S}, \subseteq\right) \longrightarrow\left(2^{X}, \subseteq\right): T \mapsto\{x \in X \mid \Gamma(x) \subseteq T\}$.
The pair $\left(\tilde{\Gamma},{ }_{*}\right)$ is an adjunction.
From this, the $\infty$-monotonicity of a Belief-function

$$
\mathrm{Bel}=P \circ *
$$

with $P$ a probability-measure follows immediately, since $P$ is $\infty$-monotone and * is meet-preserving. Furthermure it is clear, that also $\mathrm{Bel} \circ_{*}$ is $\infty$-monotone.

## Examples of adjunctions

- Lower coherent previsions:
$f: \underline{P} \mapsto \mathcal{M}(\underline{P})=\{p \in \mathscr{P}(\Omega) \mid p \geq \underline{P}\}$ and
$g: M \mapsto \underline{P}_{M}: X \mapsto \inf _{p \in M} p(X)$ are an adjunction.


## Examples of adjunctions

- Formal concept analysis:

Incidence structure $\mathbb{K}=(G, M, I)$ with $G \ldots$ objects, $M \ldots$ attributes and a relation $I \subseteq G \times M$. $(g, m) \in I$ means object $g$ has attribute $m$ (also denotad as glm).
$f:\left(2^{M}, \subseteq\right) \longrightarrow\left(2^{G}, \subseteq\right): X \mapsto\{g \in G \mid \forall m \in X: g / m\}$
",The set of all objects having all attributes in X"

$$
g:\left(2^{G}, \supseteq\right) \longrightarrow\left(2^{M}, \supseteq\right): Y \mapsto\{m \in M \mid \forall g \in Y: g / m\}
$$

"The set of all joint attributes of all objects in $Y^{\prime \prime}$.
The pair $(f, g)$ is an adjunction.

## Lemma

Let $(f, g)$ be an adjunction. Then the following holds:
A1 $g \circ f$ is extensive and $f \circ g$ is intensive.

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A3 $f \circ g \circ f=f$ and $g \circ f \circ g=g$ and thus $f \circ g$ and $g \circ f$ are idempotent．

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A6 $f$ is join-preserving and $g$ is meet-preserving.

## Lemma

- If $P$ is a complete lattice, than $f$ is a left adjoint, if and only if $f$ is join-preserving.
- If $Q$ is a complete lattice, than $g$ is a right adjoint, if and only if $g$ is meet-preserving.


## Lemma

The mapping

SER $:(\mathcal{Z}(\Omega), \leq) \longrightarrow\left(2^{B}, \subseteq\right):(X, \underline{Y}, \bar{Y}) \mapsto\{\beta \mid \mathbb{E}(\underline{Y} \mid X) \leq \beta X \leq \mathbb{E}(\bar{Y} \mid x)\}$
with

$$
\begin{aligned}
\left(X_{1}, \underline{Y}_{1}, \bar{Y}_{1}\right) \leq\left(X_{2}, \underline{Y}_{2}, \bar{Y}_{2}\right): \Longleftrightarrow & \\
& \text { i.e.: }\left(\underline{Y}_{1} \mid X\right) \geq \mathbb{E}\left(\underline{Y}_{2} \mid X\right) \& \mathbb{E}\left(\bar{Y}_{1}\right) \text { is more precise } \\
& \text { than }\left(X_{2}, \underline{Y}_{2}, \bar{Y}_{2}\right) \leq \mathbb{E}\left(\bar{Y}_{2} \mid X\right)
\end{aligned}
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is a right adjoint.

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\end{aligned}
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is a right adjoint.
The corresponding left adjoint is the ,,prediction-operator":

$$
P R: \quad\left(2^{B}, \subseteq\right) \longrightarrow(\mathcal{Z}(\Omega), \leq): M \mapsto\left(X, \min _{\beta \in M} X \beta, \max _{\beta \in M} X \beta\right)
$$

## Lemma

Thus, the following holds:
$A 1 S E R \circ P R$ is extensive and $P R \circ S E R$ is intensive.
$\mathrm{A} 2 P R$ and $S E R$ are order-preserving.
A3 $P R \circ S E R \circ P R=P R$ and $S E R \circ P R \circ S E R=S E R$ and thus $P R \circ S E R$ and $S E R \circ P R$ are idempotent.
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The mapping
$S P R:(\mathcal{Z}(\Omega), \leq) \longrightarrow\left(2^{B}, \subseteq\right):(X, \underline{Y}, \bar{Y}) \mapsto\left\{\left(X^{\prime} X\right)^{-1} X^{\prime} Y \mid \underline{Y} \leq Y \leq \bar{Y}\right\}$
is no right adjoint, since it is not meet-preserving.
In general $\operatorname{SPR}\left(Z_{1} \wedge Z_{2}\right) \neq \operatorname{SPR}\left(Z_{1}\right) \cap \operatorname{SPR}\left(Z_{2}\right)$, since the intersection of two zonoids is in general not a zonoid.
Thus, in general, only $S P R \circ P R \circ S P R \supset S P R$ holds.


## A note on sharp identification regions












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## Definition

Let $E:(P, \leq) \longrightarrow(Q, \sqsubseteq)$ be a mapping.
The monotone hull of $E$ is defined as:

$$
H(E): \quad(P, \leq) \longrightarrow(Q, \sqsubseteq): X \mapsto \bigvee_{Y \leq X} E(Y)
$$

The monotone kernel of $E$ is defined as:

$$
K(E):(P, \leq) \longrightarrow(Q, \sqsubseteq): X \mapsto \bigwedge_{Y \geq X} E(Y) .
$$

These set-valued mappings are both order-preserving (i.e: $X \leq Y \Longrightarrow(H(E))(X) \sqsubseteq(H(E))(Y) \quad \& \quad(K(E))(X) \sqsubseteq(K(E))(Y))$.

## A criterion-function-based mapping

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## Lemma

Let the criterion-function $Q: B \longrightarrow \mathbb{R}$ be defined as:

$$
Q(\beta)=\int\left\{(\mathbb{E}(\underline{Y} \mid x)-x \beta)_{+}^{2}+(\mathbb{E}(\bar{Y} \mid x)-x \beta)_{-}^{2}\right\} d \mathbb{P}(x) .
$$

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Then the criterion-based mapping

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E_{Q}: \mathcal{Z}(\Omega) \longrightarrow 2^{B}:(X, \underline{Y}, \bar{Y}) \mapsto \underset{\beta \in B}{\operatorname{argmin}} Q(\beta)
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## A criterion-function-based mapping

## Lemma

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\begin{aligned}
S P R & =H\left(E_{Q}\right) \\
S E R & =K\left(E_{Q}\right) .
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## Estimation of SER and SPR

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In general, there is no consistent and (in a certain sense) robust estimator of SER.

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