# A note on sharp identification regions

Let  $P := \{ \mathbb{P}_{\theta} \mid \theta \in \Theta \}$  be a statistical model and

- Y,... unobservable random variables,
- $X, \underline{Y}, \overline{Y}, \ldots$  observable random variables w.r.t an underlying probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ .
- The joint distribution of the random Variables  $X, Y, \underline{Y}, \overline{Y}$  under a model  $P_{\theta}$  is denoted with  $F_{\theta}$  and the joint distribution under the "true model"  $\mathbb{P}$  is denoted with  $F^{X,Y,\underline{Y},\overline{Y}}$ .
- The unobserved variables fullfill a certain condition  $C(X, Y, \underline{Y}, \overline{Y}) = 1$ .

e.g. 
$$\underline{Y} \leq Y \leq \overline{Y}$$
 or  $\forall X : \mathbb{E}(\underline{Y} \mid X) \leq \mathbb{E}(Y \mid X) \leq \mathbb{E}(\overline{Y} \mid X)$ .



• Two parameters  $\theta_1$  and  $\theta_2$  are undistinguishable (i.e.  $\theta_1 \sim \theta_2$ ) if the corresponding models  $\mathbb{P}_{\theta_1}$  and  $\mathbb{P}_{\theta_2}$  are empirically undistinguishable, which means, that the distributions of the observable variables are the same:

$$F_{\theta_{\mathbf{1}}}^{X,\underline{Y},\overline{Y}}=F_{\theta_{\mathbf{2}}}^{X,\underline{Y},\overline{Y}}.$$



A statistical model P is called point-identified, if any two different parameters  $\theta_1$  and  $\theta_2$  are empirically distinguishable, i.e.:

$$\sim \ = \ \Delta_{\Theta} = \{(\theta, \theta) \mid \theta \in \Theta\}.$$

Otherwise it is called partially identified.

### Example

The simple linear model

$$\Theta = B imes \mathbb{R}_{\geq 0} imes \mathcal{Z}(\mathbb{R}_{\geq 0}) imes \mathcal{Z}(\mathbb{R}_{\geq 0})$$

with  $B = \mathbb{R}^2$ . For  $\theta = (\beta, \sigma^2, \sigma_I, \sigma_u) \in \Theta$ , the random variables are defined as:

$$\begin{array}{rcl} Y & = & X\beta + \varepsilon \\ \underline{Y} & = & X\beta + \varepsilon - \sigma_I \\ \overline{Y} & = & X\beta + \varepsilon + \sigma_u \end{array}$$

with  $\varepsilon \sim N(0, \sigma^2 I)$ .



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with  $\varepsilon \sim N(0, \sigma^2 I)$ .

Here we are only interested in the values of  $\beta \in B$ .



This model is only partially identified. For example

$$((\beta_0, \beta_1), \sigma^2, 0, 1)$$
  $\sim ((\beta_0 + 1, \beta_1), \sigma^2, 1, 0).$ 

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Moreover, the quotient space  $\Theta_{/\sim}$  ist not of the form

$$\Theta_{/\sim} = B_{/\approx} \times \text{,,rest''},$$

so we must factorize the whole space  $\Theta$  and not only the interesting B to make the model point-identified.

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Estimation  $-- \mid \mathit{Model} \mid \longrightarrow \mid \mathit{Pediction} \mid$ ..model as a truth to be estimated " "model as a tool to be applied" e.g.: least squares estimator linear model best linear predictor

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$$\begin{array}{ccc} & Y \sim F_{\theta} \\ \Longleftrightarrow & F^{Y} = F_{\theta}^{Y} \end{array}$$

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$$\iff L(F_{\theta}^{Y}, F) = 0$$

for some distance-function  $L(\cdot, \cdot)$ .



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The actual problem is, that  $F^Y$  is unknown  $\Longrightarrow$  later.



Let  $P = \{ \mathbb{P}_{\theta} \mid \theta \in \Theta \}$  be a statistical model with the corresponding joint distributions  $\{ F_{\theta}^{X,Y,\underline{Y},\overline{Y}} \mid \theta \in \Theta \}$  and  $X,\underline{Y},\overline{Y}$  random variables with the joint distribution  $F^{X,\underline{Y},\overline{Y}}$ . The **Sharp Estimation Region (SER)** is defined as:



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$$SER(\underline{Y}, \overline{Y}) = \underset{\theta \in \Theta}{\mathsf{argmin}} \left( \inf_{Y \leq t. C(X, Y, \underline{Y}, \overline{Y}) = 1} L\left(F_{\theta}, F^{X, Y, \overline{Y}, \underline{Y}}\right) \right).$$



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### Now: Linear Model

We are only interested in the components  $(\beta_0, \beta_1)$  of an element  $\theta = ((\beta_0, \beta_1), \sigma^2, \sigma_l, \sigma_u) \in SER$  and denote the set

$$\{(\beta_0,\beta_1) \mid ((\beta_0,\beta_1),\sigma^2,\sigma_I,\sigma_u) \in SER\}$$

as the sharp estimation region (analogously for the sharp prediction region).

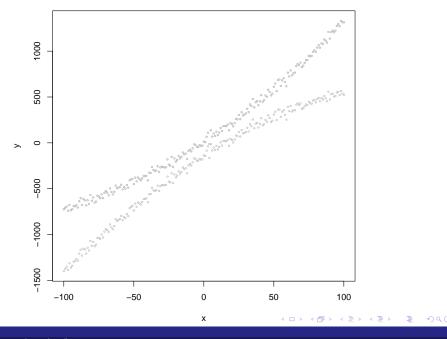


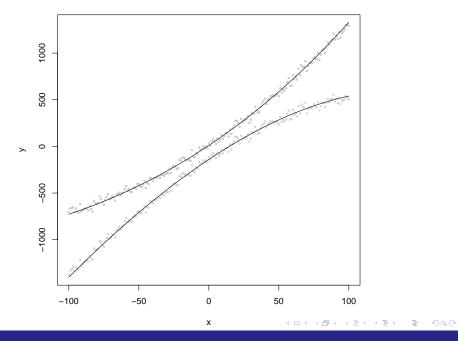
#### Linear Model

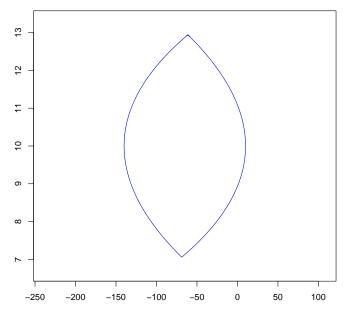
$$SER = \{ \beta \in B | \mathbb{E}(\underline{Y} \mid X) \le X\beta \le \mathbb{E}(\overline{Y} \mid X) \}$$

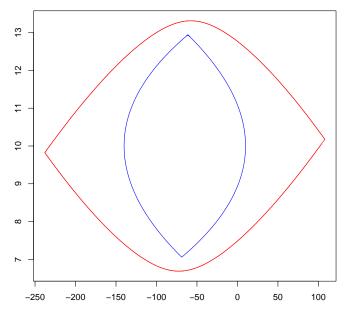
$$SPR = \{ \underset{\beta \in B}{\operatorname{argmin}} \mathbb{E}((X\beta - Y)^2) \mid Y \in [\underline{Y}, \overline{Y}] \}$$

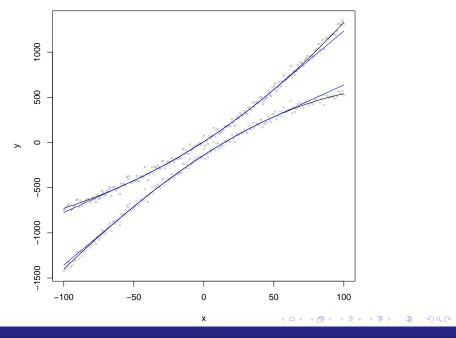
$$= \{ (X'X)^{-1}X'Y \mid Y \in [\underline{Y}, \overline{Y}] \}$$

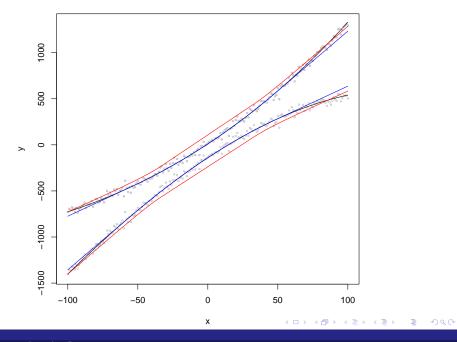












### **Theorem**

Let  $I \subset \mathbb{R}^2$  be a compact convex set. Then there exist random variables  $X, Y, \overline{Y}$  such that

$$SER(X, \underline{Y}, \overline{Y}) = I,$$

namely:

$$\begin{array}{lcl} X & \sim & \textit{N}(0,1) \\ \underline{Y} & = & \min\{\beta_0 + \beta_1 X \mid (\beta_0,\beta_1) \in \textit{I}\} \\ \underline{Y} & = & \max\{\beta_0 + \beta_1 X \mid (\beta_0,\beta_1) \in \textit{I}\}. \end{array}$$

The Minkowski-Sum

$$M = \bigoplus_{i=1}^{n} l_i = \left\{ \sum_{i=1}^{n} p_i \mid p_i \in l_i \right\}$$

of n line-segments  $l_i \subseteq \mathbb{R}^d$  is called a **zonotope**.

A zonotope is a convex, compact and centrally symmetric polytope with finite many extremepoints and central-symmetric facets.

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#### Definition

A closed, centrally symmetric convex set  $Z \subseteq \mathbb{R}^d$  is called a **zonoid**, if it can be approximated arbitrarily closely by zonotopes (w.r.t. a metric, e.g. the Hausdorff distance).

For d=2 the zonoids are exactly the closed, centrally symmetric convex sets.



Let  $I \subseteq \mathbb{R}^2$  be a zonoid in general position. Then there exists random variables  $X,\underline{Y},\overline{Y}$  such that

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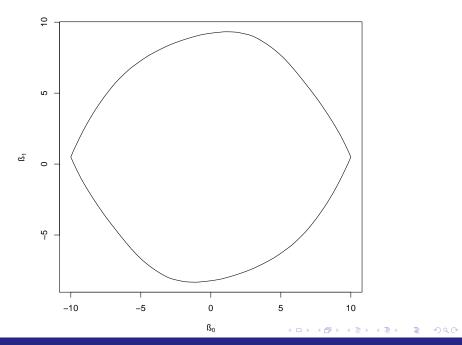
### Lemma

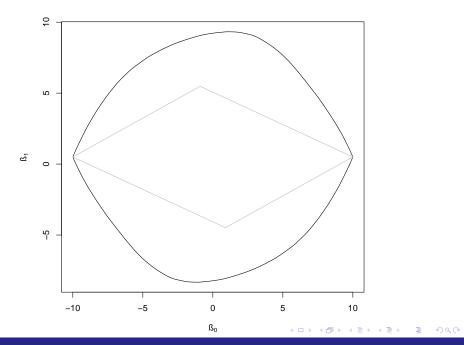
Let  $I = SPR(X, \underline{Y}^*, \overline{Y}^*) \subseteq \mathbb{R}^2$  be a zonoid and  $E \subseteq SER(X, \underline{Y}^*, \overline{Y}^*)$  an arbitrary compact convex set. Then for every  $\varepsilon > 0$  there exist random variables  $X, \underline{Y}, \overline{Y}$  such that:

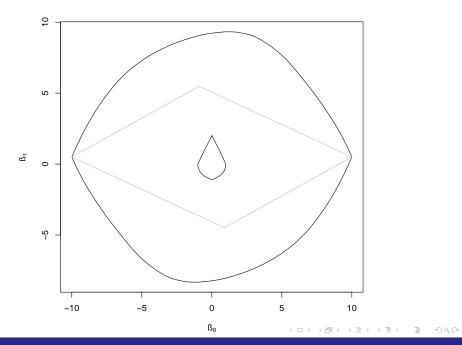
$$d_H(SPR(X, \underline{Y}, \overline{Y}), I) \leq \varepsilon$$
  
 $d_H(SER(X, \underline{Y}, \overline{Y}), E) \leq \varepsilon$ 

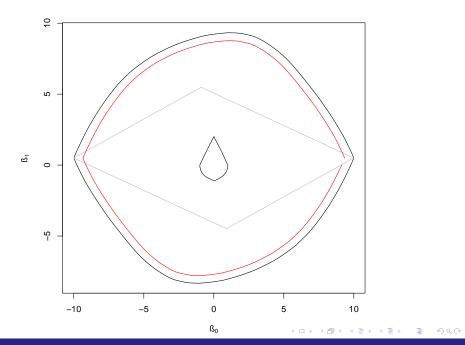
with the Hausdorff distance  $d_H$ .

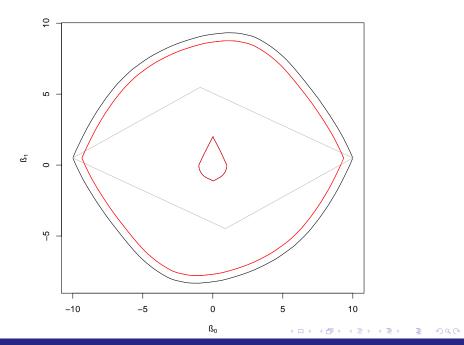












# Mappings between ordered sets

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## Definition

Let  $(P, \leq)$  and  $(Q, \sqsubseteq)$  be partially ordered sets. A pair (f, g) of mappings  $f: P \longrightarrow Q$  and  $g: Q \longrightarrow P$  is called **adjunction**, if:

$$\forall p \in P \forall q \in Q: p \leq g(q) \iff f(p) \sqsubseteq q.$$

In this case, f is called **left adjoint** and g is called **right adjoint**.



• Dempster-Shafer-Theory:  $Multivalued\ mapping\ \Gamma: X \longrightarrow 2^S\ with\ corresponding$ 

$$\tilde{\Gamma}:(2^X,\subseteq)\longrightarrow(2^S,\subseteq):A\mapsto\bigcup_{a\in A}\Gamma(a)$$
 and the operator

$$_*: (2^S, \subseteq) \longrightarrow (2^X, \subseteq): T \mapsto \{x \in X \mid \Gamma(x) \subseteq T\}.$$

The pair  $(\tilde{\Gamma},\ _*)$  is an adjunction. From this, the  $\infty$ -monotonicity of a Belief-function

$$Bel = P \circ *$$

with P a probability-measure follows immediately, since P is  $\infty$ -monotone and \* is meet-preserving. Furthermure it is clear, that also Bel  $\circ$  \* is  $\infty$ -monotone.



Lower coherent previsions:

$$f: \underline{P} \mapsto \mathcal{M}(\underline{P}) = \{p \in \mathscr{P}(\Omega) \mid p \geq \underline{P}\}$$
 and

 $g: M \mapsto \underline{P}_M: X \mapsto \inf_{p \in M} p(X)$  are an adjunction.



• Formal concept analysis: Incidence structure  $\mathbb{K} = (G, M, I)$ with  $G \dots$  objects,  $M \dots$  attributes and a relation  $I \subseteq G \times M$ .  $(g, m) \in I$  means object g has attribute m (also denoted as glm).

$$f: (2^{M}, \subseteq) \longrightarrow (2^{G}, \subseteq): X \mapsto \{g \in G | \forall m \in X : glm\}$$

"The set of all objects having all attributes in X"

$$g:(2^G,\supseteq)\longrightarrow (2^M,\supseteq):Y\mapsto \{m\in M\mid \forall g\in Y:glm\}$$

"The set of all joint attributes of all objects in Y".

The pair (f,g) is an adjunction.



Let (f,g) be an adjunction. Then the following holds:

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- A6 f is join-preserving and g is meet-preserving.



- If P is a complete lattice, than f is a left adjoint, if and only if f is join-preserving.
- If Q is a complete lattice, than g is a right adjoint, if and only if g is meet-preserving.

The mapping

$$\textit{SER} \quad : \quad (\mathcal{Z}(\Omega), \leq) \longrightarrow (2^{\textit{B}}, \subseteq) : (X, \underline{Y}, \overline{Y}) \mapsto \{\beta \mid \mathbb{E}(\underline{Y} \mid X) \leq \beta X \leq \mathbb{E}(\overline{Y} \mid x)\}$$

with

$$\boxed{ (X_1, \underline{Y_1}, \overline{Y_1}) \leq (X_2, \underline{Y_2}, \overline{Y_2}) :} \iff \boxed{ \boxed{ \mathbb{E}(\underline{Y_1} \mid X) \geq \mathbb{E}(\underline{Y_2} \mid X) \ \& \ \mathbb{E}(\overline{Y_1} \mid X) \leq \mathbb{E}(\overline{Y_2} \mid X) } }$$
 
$$\textit{i.e.:} \ (X_1, \underline{Y_1}, \overline{Y_1}) \textit{ is more precise}$$
 
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$$\frac{\left[ (X_{1}, \underline{Y}_{1}, \overline{Y}_{1}) \leq (X_{2}, \underline{Y}_{2}, \overline{Y}_{2}) : \right]}{\text{i.e.: } \left( X_{1}, \underline{Y}_{1}, \overline{Y}_{1} \right) \text{ is more precise} }$$

$$\text{than } \left( X_{2}, \underline{Y}_{2}, \overline{Y}_{2} \right)$$

is a right adjoint.

The corresponding left adjoint is the "prediction-operator":

$$PR: \qquad (2^B, \subseteq) \longrightarrow \left(\mathcal{Z}(\Omega), \leq): M \mapsto (X, \min_{\beta \in M} X\beta, \max_{\beta \in M} X\beta\right).$$

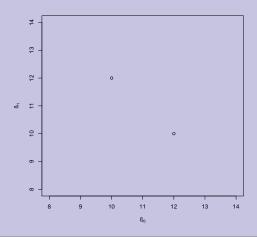


Thus, the following holds:

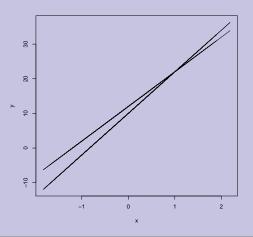
- A1 SER o PR is extensive and PR o SER is intensive.
- A2 PR and SER are order-preserving.
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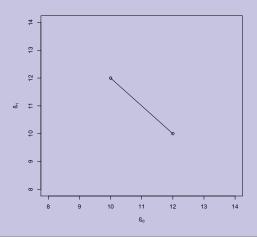


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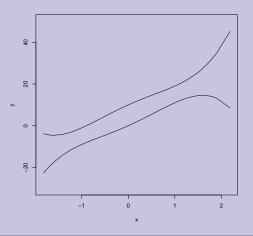


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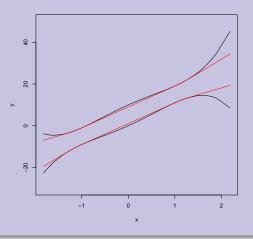


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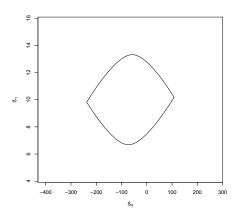
The mapping

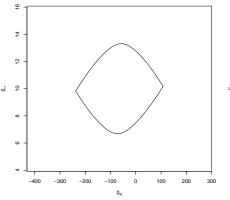
$$\textit{SPR}: (\mathcal{Z}(\Omega), \leq) \longrightarrow (2^{\textit{B}}, \subseteq): (X, \underline{Y}, \overline{Y}) \mapsto \{(X'X)^{-1}X'Y \mid \underline{Y} \leq Y \leq \overline{Y}\}$$

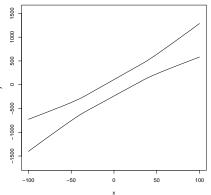
is no right adjoint, since it is not meet-preserving.

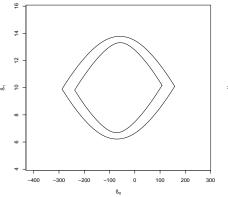
In general  $SPR(Z_1 \land Z_2) \neq SPR(Z_1) \cap SPR(Z_2)$ , since the intersection of two zonoids is in general not a zonoid.

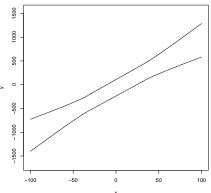
Thus, in general, only SPR  $\circ$  PR  $\circ$  SPR  $\supset$  SPR holds.

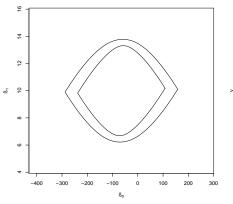


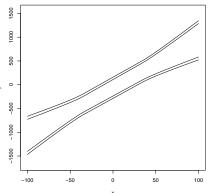


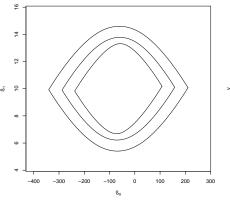


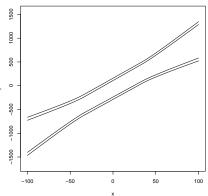


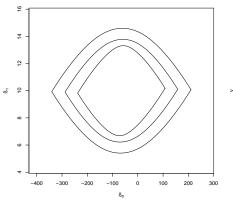


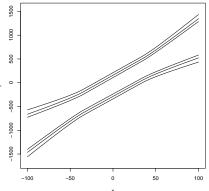


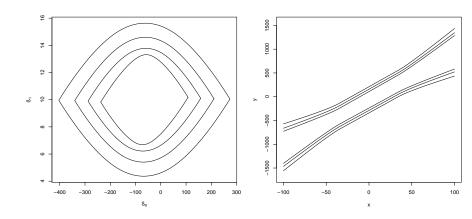


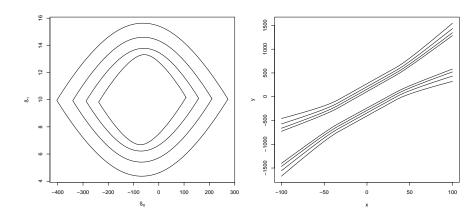












## Definition

Let  $E:(P,\leq)\longrightarrow (Q,\sqsubseteq)$  be a mapping.

The monotone hull of E is defined as:

$$H(E)$$
 :  $(P, \leq) \longrightarrow (Q, \sqsubseteq) : X \mapsto \bigvee_{Y \leq X} E(Y)$ .

The monotone kernel of E is defined as:

$$K(E)$$
:  $(P, \leq) \longrightarrow (Q, \sqsubseteq) : X \mapsto \bigwedge_{Y>X} E(Y).$ 

These set-valued mappings are both order-preserving

$$(\textit{i.e.} \ X \leq Y \Longrightarrow (\textit{H}(\textit{E}))(X) \sqsubseteq (\textit{H}(\textit{E}))(Y) \quad \& \quad (\textit{K}(\textit{E}))(X) \sqsubseteq (\textit{K}(\textit{E}))(Y)).$$

### Lemma

Let the criterion-function  $Q: B \longrightarrow \mathbb{R}$  be defined as:

$$Q(\beta) = \int \left\{ \left( \mathbb{E}(\underline{Y}|x) - x\beta \right)_{+}^{2} + \left( \mathbb{E}(\overline{Y}|x) - x\beta \right)_{-}^{2} \right\} d\mathbb{P}(x).$$

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Then the criterion-based mapping

$$E_Q: \mathcal{Z}(\Omega) \longrightarrow 2^B: (X, \underline{Y}, \overline{Y}) \mapsto \operatorname*{argmin}_{\beta \in B} Q(\beta)$$

is a source of SER and SPR:



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$$SPR = H(E_Q)$$
  
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## Estimation of SER and SPR



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In general, there is no monotone, nonpartial, consistent estimator of SER.

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In general, there is no consistent and (in a certain sense) robust estimator of SER.

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