

The course of well-being over the life span - Restricted Likelihood Ratio Testing (RLRT) in the presence of correlated errors

Andrea Wiencierz in cooperation with:
Sonja Greven and Helmut Küchenhoff

Department for Statistics, LMU Munich

June 30, 2010

Contents

- 1 Introduction
- 2 RLRT for mixed model smoothing
 - General form linear mixed models
 - Restricted Likelihood Ratio Test
- 3 GLS transformation
- 4 RLRT with GLS transformation
 - Simulations
 - Application to SOEP data
- 5 Open questions

- 1 Introduction
- 2 RLRT for mixed model smoothing
 - General form linear mixed models
 - Restricted Likelihood Ratio Test
- 3 GLS transformation
- 4 RLRT with GLS transformation
 - Simulations
 - Application to SOEP data
- 5 Open questions

“Happiness economics”

Question of interest:

“How does (subjective) well-being change over the life cycle?”

“Happiness economics”

Question of interest:

“How does (subjective) well-being change over the life cycle?”

Previous work:

- Consulting project *“Well-being and Age II”*: several parametric analyses of SOEP data and construction of a semiparametric model,
- Discussion paper *“Well-being over the life span: semiparametric evidence from British and German longitudinal data”*: application of semiparametric models to BHPS and SOEP data.

“Happiness economics”

Question of interest:

“How does (subjective) well-being change over the life cycle?”

Previous work:

- Consulting project *“Well-being and Age II”*: several parametric analyses of SOEP data and construction of a semiparametric model,
- Discussion paper *“Well-being over the life span: semiparametric evidence from British and German longitudinal data”*: application of semiparametric models to BHPS and SOEP data.

⇒ Lack of a valid test to confirm the findings.

- 1 Introduction
- 2 RLRT for mixed model smoothing
 - General form linear mixed models
 - Restricted Likelihood Ratio Test
- 3 GLS transformation
- 4 RLRT with GLS transformation
 - Simulations
 - Application to SOEP data
- 5 Open questions

Linear mixed model (LMM)

Simple linear mixed model for longitudinal data of individuals $i = 1, \dots, n$ with n_i observations:

$$\mathbf{y}_i = \mathbf{X}_i \boldsymbol{\beta} + \mathbf{Z}_i \mathbf{b}_i + \boldsymbol{\epsilon}_i,$$

$$\mathbf{b}_i \sim N(\mathbf{0}, \tilde{\mathbf{D}}), \quad \text{where } \tilde{\mathbf{D}} = \sigma_b \mathbf{I}_K,$$

$$\boldsymbol{\epsilon}_i \sim N(\mathbf{0}, \mathbf{R}_i), \quad \text{where } \mathbf{R}_i = \sigma_\epsilon \mathbf{I}_{n_i},$$

$\mathbf{b}_i, \mathbf{b}_j$ and $\boldsymbol{\epsilon}_i, \boldsymbol{\epsilon}_j$ independent for all $i \neq j$,

$\mathbf{b}_i, \boldsymbol{\epsilon}_j$ independent for all i, j .

Linear mixed model (LMM)

Simple linear mixed model for longitudinal data of individuals $i = 1, \dots, n$ with n_i observations:

$$\mathbf{y}_i = \mathbf{X}_i \boldsymbol{\beta} + \mathbf{Z}_i \mathbf{b}_i + \boldsymbol{\epsilon}_i,$$

$$\mathbf{b}_i \sim N(\mathbf{0}, \tilde{\mathbf{D}}), \quad \text{where } \tilde{\mathbf{D}} = \sigma_b \mathbf{I}_K,$$

$$\boldsymbol{\epsilon}_i \sim N(\mathbf{0}, \mathbf{R}_i), \quad \text{where } \mathbf{R}_i = \sigma_\epsilon \mathbf{I}_{n_i},$$

$\mathbf{b}_i, \mathbf{b}_j$ and $\boldsymbol{\epsilon}_i, \boldsymbol{\epsilon}_j$ independent for all $i \neq j$,

$\mathbf{b}_i, \boldsymbol{\epsilon}_j$ independent for all i, j .

Corresponding global model:

$$\mathbf{y} = \mathbf{X} \boldsymbol{\beta} + \mathbf{Z} \mathbf{b} + \boldsymbol{\epsilon},$$

$$\mathbf{b} \sim N(\mathbf{0}, \mathbf{D}), \quad \text{where } \mathbf{D} = \text{diag}(\tilde{\mathbf{D}}) = \sigma_b \mathbf{I}_{nK},$$

$$\boldsymbol{\epsilon} \sim N(\mathbf{0}, \mathbf{G}), \quad \text{where } \mathbf{G} = \text{diag}(\mathbf{G}_i) = \sigma_\epsilon \mathbf{I}_N,$$

$\mathbf{b}, \boldsymbol{\epsilon}$ independent.

General form linear mixed model

General formulation of a mixed model for $i = 1, \dots, N$ data points:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}_1 \mathbf{b}_1 + \dots + \mathbf{Z}_S \mathbf{b}_S + \boldsymbol{\epsilon},$$

$$\mathbf{b}_s \sim \mathcal{N}(\mathbf{0}, \sigma_s \boldsymbol{\Sigma}_s), \quad s = 1, \dots, S,$$

$$\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \sigma_\epsilon \mathbf{I}_N),$$

$\mathbf{b}_1, \dots, \mathbf{b}_S, \boldsymbol{\epsilon}$ independent.

General form linear mixed model

General formulation of a mixed model for $i = 1, \dots, N$ data points:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}_1 \mathbf{b}_1 + \dots + \mathbf{Z}_S \mathbf{b}_S + \boldsymbol{\epsilon},$$

$$\mathbf{b}_s \sim \mathcal{N}(\mathbf{0}, \sigma_s \boldsymbol{\Sigma}_s), \quad s = 1, \dots, S,$$

$$\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \sigma_\epsilon \mathbf{I}_N),$$

$\mathbf{b}_1, \dots, \mathbf{b}_S, \boldsymbol{\epsilon}$ independent.

Nonparametric model (with cubic TP basis) as a mixed model:

$$y_i = f(x_i) + \epsilon_i = \sum_{j=1}^K \gamma_j B_j(x_i) + \epsilon_i$$

General form linear mixed model

General formulation of a mixed model for $i = 1, \dots, N$ data points:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}_1 \mathbf{b}_1 + \dots + \mathbf{Z}_S \mathbf{b}_S + \boldsymbol{\epsilon},$$

$$\mathbf{b}_s \sim \mathcal{N}(\mathbf{0}, \sigma_s \boldsymbol{\Sigma}_s), \quad s = 1, \dots, S,$$

$$\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \sigma_\epsilon \mathbf{I}_N),$$

$\mathbf{b}_1, \dots, \mathbf{b}_S, \boldsymbol{\epsilon}$ independent.

Nonparametric model (with cubic TP basis) as a mixed model:

$$\begin{aligned} y_i &= f(x_i) + \epsilon_i = \sum_{j=1}^K \gamma_j B_j(x_i) + \epsilon_i \\ &= \underbrace{\gamma_0 + \gamma_1 x_i + \gamma_2 x_i^2 + \gamma_3 x_i^3}_{=: \mathbf{x}_i \boldsymbol{\beta}} + \underbrace{\sum_{j=1}^K \gamma_{j+3} (x_i - \kappa_j)_+^3}_{=: \mathbf{z}_i \mathbf{b}} + \epsilon_i \end{aligned}$$

- 1 Introduction
- 2 RLRT for mixed model smoothing
 - General form linear mixed models
 - Restricted Likelihood Ratio Test
- 3 GLS transformation
- 4 RLRT with GLS transformation
 - Simulations
 - Application to SOEP data
- 5 Open questions

Testing for zero variance components

We want to test:

$$H_0 : b_1 = \dots = b_K = 0 \quad \text{vs.} \quad H_1 : b_j \neq 0 \text{ for one } j.$$

This is equivalent to testing:

$$H_0 : \sigma_b = 0 \quad \text{vs.} \quad H_1 : \sigma_b > 0.$$

Testing for zero variance components

We want to test:

$$H_0 : b_1 = \dots = b_K = 0 \quad \text{vs.} \quad H_1 : b_j \neq 0 \text{ for one } j.$$

This is equivalent to testing:

$$H_0 : \sigma_b = 0 \quad \text{vs.} \quad H_1 : \sigma_b > 0.$$

Violation of assumptions of usual asymptotics for Likelihood Ratio tests:

- the tested parameter is on the boundary of the parameter space,
- observations are not i.i.d.

Testing for zero variance components

We want to test:

$$H_0 : b_1 = \dots = b_K = 0 \quad \text{vs.} \quad H_1 : b_j \neq 0 \text{ for one } j.$$

This is equivalent to testing:

$$H_0 : \sigma_b = 0 \quad \text{vs.} \quad H_1 : \sigma_b > 0.$$

Violation of assumptions of usual asymptotics for Likelihood Ratio tests:

- the tested parameter is on the boundary of the parameter space,
- observations are not i.i.d.

⇒ Test distribution?

RLRT for zero variance components

Exact finite sample null distribution of the RLRT (Crainiceanu and Ruppert, 2004):

$$RLRT_n \stackrel{d}{=} \sup_{\lambda \geq 0} \left\{ (n-p) \log \left[1 + \frac{N_n(\lambda)}{D_n(\lambda)} \right] - \sum_{l=1}^K \log(1 + \lambda \mu_{l,n}) \right\},$$

where

$$N_n(\lambda) = \sum_{l=1}^K \frac{\lambda \mu_{l,n}}{1 + \lambda \mu_{l,n}} \omega_l^2, \quad D_n(\lambda) = \sum_{l=1}^K \frac{\omega_l^2}{1 + \lambda \mu_{l,n}} + \sum_{l=K+1}^{n-p} \omega_l^2,$$

$\mu_{l,n}$ eigenvalues of $\Sigma^{\frac{1}{2}} \mathbf{Z}' (\mathbf{I}_n - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}') \mathbf{Z} \Sigma^{\frac{1}{2}}$ and $\omega_l \stackrel{iid}{\sim} N(0, 1)$.

RLRT for zero variance components

Exact finite sample null distribution of the RLRT (Crainiceanu and Ruppert, 2004):

$$RLRT_n \stackrel{d}{=} \sup_{\lambda \geq 0} \left\{ (n-p) \log \left[1 + \frac{N_n(\lambda)}{D_n(\lambda)} \right] - \sum_{l=1}^K \log(1 + \lambda \mu_{l,n}) \right\},$$

where

$$N_n(\lambda) = \sum_{l=1}^K \frac{\lambda \mu_{l,n}}{1 + \lambda \mu_{l,n}} \omega_l^2, \quad D_n(\lambda) = \sum_{l=1}^K \frac{\omega_l^2}{1 + \lambda \mu_{l,n}} + \sum_{l=K+1}^{n-p} \omega_l^2,$$

$\mu_{l,n}$ eigenvalues of $\Sigma^{\frac{1}{2}} \mathbf{Z}' (\mathbf{I}_n - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}') \mathbf{Z} \Sigma^{\frac{1}{2}}$ and $\omega_l \stackrel{iid}{\sim} N(0, 1)$.

⇒ This distribution can be simulated in R with the RLRsim package (Scheipl, 2010).

- 1 Introduction
- 2 RLRT for mixed model smoothing
 - General form linear mixed models
 - Restricted Likelihood Ratio Test
- 3 GLS transformation
- 4 RLRT with GLS transformation
 - Simulations
 - Application to SOEP data
- 5 Open questions

The idea of GLS transformation

Consider the simple linear model:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}, \quad \text{where } \boldsymbol{\epsilon} \sim N(\mathbf{0}, \mathbf{V}) \text{ and } \mathbf{V} \text{ is known.}$$

Then the generalized least squares estimator is efficient:

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$$

The idea of GLS transformation

Consider the simple linear model:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}, \quad \text{where } \boldsymbol{\epsilon} \sim \mathbf{N}(\mathbf{0}, \mathbf{V}) \text{ and } \mathbf{V} \text{ is known.}$$

Then the generalized least squares estimator is efficient:

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y} = (\tilde{\mathbf{X}}'\tilde{\mathbf{X}})^{-1}\tilde{\mathbf{X}}'\tilde{\mathbf{y}},$$

where $\tilde{\mathbf{X}} := \mathbf{V}^{-\frac{1}{2}}\mathbf{X}$ and $\tilde{\mathbf{y}} := \mathbf{V}^{-\frac{1}{2}}\mathbf{y}$.

The idea is to estimate \mathbf{V} , then transform the data and estimate $\hat{\boldsymbol{\beta}}$ from the transformed data.

The idea of GLS transformation

Consider the simple linear model:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}, \quad \text{where } \boldsymbol{\epsilon} \sim \mathbf{N}(\mathbf{0}, \mathbf{V}) \text{ and } \mathbf{V} \text{ is known.}$$

Then the generalized least squares estimator is efficient:

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y} = (\tilde{\mathbf{X}}'\tilde{\mathbf{X}})^{-1}\tilde{\mathbf{X}}'\tilde{\mathbf{y}},$$

where $\tilde{\mathbf{X}} := \mathbf{V}^{-\frac{1}{2}}\mathbf{X}$ and $\tilde{\mathbf{y}} := \mathbf{V}^{-\frac{1}{2}}\mathbf{y}$.

The idea is to estimate \mathbf{V} , then transform the data and estimate $\hat{\boldsymbol{\beta}}$ from the transformed data.

In the case of the general form LMM:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}_1\mathbf{b}_1 + \dots + \mathbf{Z}_S\mathbf{b}_S + \boldsymbol{\epsilon}, \quad \text{where } \boldsymbol{\epsilon} \sim \mathbf{N}(\mathbf{0}, \sigma_\epsilon \mathbf{R})$$

the transformation leads to the model:

$$\tilde{\mathbf{y}} = \tilde{\mathbf{X}}\boldsymbol{\beta} + \tilde{\mathbf{Z}}_1\mathbf{b}_1 + \dots + \tilde{\mathbf{Z}}_S\mathbf{b}_S + \tilde{\boldsymbol{\epsilon}}, \quad \text{where } \tilde{\boldsymbol{\epsilon}} \sim \mathbf{N}(\mathbf{0}, \sigma_\epsilon \mathbf{I}_N)$$

- 1 Introduction
- 2 RLRT for mixed model smoothing
 - General form linear mixed models
 - Restricted Likelihood Ratio Test
- 3 GLS transformation
- 4 RLRT with GLS transformation
 - Simulations
 - Application to SOEP data
- 5 Open questions

Scenarios

1 **kub_testRI**

H_1 model: simple cubic LMM with one random intercept and autocorrelated errors,

test: RLRT for zero variance of the random intercepts

2 **kubglatt**

H_1 model: smoothing LMM (TP basis 3rd degree) with autocorrelated errors,

test: RLRT for zero variance of the penalized smoothing coefficients

3 **kubglatt_RI**

H_1 model: smoothing LMM (TP basis 3rd degree) with a nuisance random intercept and autocorrelated errors,

test: RLRT for zero variance of the penalized smoothing coefficients

All scenarios simulated for $I = 20, 100, 500$ individuals with $J = 4, 10, 20$ observations and for $\rho = 0, 0.4, 0.8$.

Scenarios

1 **kub_testRI**

H_1 model: simple cubic LMM with one random intercept and autocorrelated errors,

test: RLRT for zero variance of the random intercepts

2 **kubglatt**

H_1 model: smoothing LMM (TP basis 3rd degree) with autocorrelated errors,

test: RLRT for zero variance of the penalized smoothing coefficients

3 **kubglatt_RI**

H_1 model: smoothing LMM (TP basis 3rd degree) with a nuisance random intercept and autocorrelated errors,

test: RLRT for zero variance of the penalized smoothing coefficients

All scenarios simulated for $I = 20, 100, 500$ individuals with $J = 4, 10, 20$ observations and for $\rho = 0, 0.4, 0.8$.

⇒ All type 1 error rates in $[0.045, 0.06]$, no systematic variation observed.

- 1 Introduction
- 2 RLRT for mixed model smoothing
 - General form linear mixed models
 - Restricted Likelihood Ratio Test
- 3 GLS transformation
- 4 RLRT with GLS transformation
 - Simulations
 - Application to SOEP data
- 5 Open questions

SOEP data:

- 253 044 observations of 33 451 individuals
- observations from 20 waves between 1986 to 2007
- well-being measured by self-reported overall life satisfaction on a discrete scale from 0 to 10

SOEP data:

- 253 044 observations of 33 451 individuals
- observations from 20 waves between 1986 to 2007
- well-being measured by self-reported overall life satisfaction on a discrete scale from 0 to 10

Considered test:

- H_0 : LMM with a cubic polynomial in *age* with nuisance random intercept and autocorrelated errors vs.
- H_1 : Semiparametric LMM with nuisance random intercept and autocorrelated errors.

SOEP data:

- 253 044 observations of 33 451 individuals
- observations from 20 waves between 1986 to 2007
- well-being measured by self-reported overall life satisfaction on a discrete scale from 0 to 10

Considered test:

- H_0 : LMM with a cubic polynomial in *age* with nuisance random intercept and autocorrelated errors vs.
- H_1 : Semiparametric LMM with nuisance random intercept and autocorrelated errors.

Result:

- observed RLRT: 557.0264
- 95% quantile of the approximate distribution: 1.612736

⇒ H_0 is rejected.

- 1 Introduction
- 2 RLRT for mixed model smoothing
 - General form linear mixed models
 - Restricted Likelihood Ratio Test
- 3 GLS transformation
- 4 RLRT with GLS transformation
 - Simulations
 - Application to SOEP data
- 5 Open questions

Open Issues

- Investigate different strategies/ methods to estimate the correlation structure and their potentials/ problems.
- Quantify the approximation bias of GLS transformation.
- Exact finite sample or asymptotic distribution of RLRT in case of more than one random effects variance, instead of well-proven approximation?
- Theoretical derivation of the distribution of RLRT without GLS transformation?
- ...

References

- Crainiceanu, C. and D. Ruppert (2004). Likelihood ratio tests in linear mixed models with one variance component. *Journal of the Royal Statistical Society. Series B (Statistical Methodology)* 66(1), 165–185.
- Greven, S. (2007). *Non-Standard Problems in Inference for Additive and Linear Mixed Models*. Ph. D. thesis, LMU Munich.
- Scheipl, F. (2010). *RLRsim: Exact (Restricted) Likelihood Ratio tests for mixed and additive models*. R package version 2.0-5.
- Wunder, C., A. Wiencierz, J. Schwarze, H. Küchenhoff, S. Kleyer, and P. Bleninger (2009). Well-being over the life span: Semiparametric evidence from british and german longitudinal data. Technical Report 179, DIW Berlin.