

# "Strong happiness" and other properties of certain imprecise probability models when treating samples sequentially

Gero Walter

Institut für Statistik  
Ludwig-Maximilians-Universität München

May 12th, 2010



Institut  
für  
Statistik





# Prior-Data Conflict

Prior-Data Conflict  $\hat{=}$  situation in which...

- ▶ ... *informative prior beliefs* and *trusted data* (sampling model correct, no outliers, etc.) are in conflict
- ▶ "... the prior [places] its mass primarily on distributions in the sampling model for which the observed data is surprising." (Evans & Moshonov, 2006)



## Simple Example: Dirichlet-Multinomial-Model

Data:	$\mathbf{k}$	$\sim$	$M(\boldsymbol{\theta})$	$(\sum k_j = n)$
conjugate prior:	$\boldsymbol{\theta}$	$\sim$	$\text{Dir}(\boldsymbol{\alpha})$	$(\sum \theta_j = 1)$
posterior:	$\boldsymbol{\theta} \mid \mathbf{k}$	$\sim$	$\text{Dir}(\boldsymbol{\alpha} + \mathbf{k})$	

$$\mathbb{E}[\theta_j] = \frac{\alpha_j}{\sum \alpha_i}$$

$$\mathbb{V}(\theta_j) = \frac{\mathbb{E}[\theta_j](1 - \mathbb{E}[\theta_j])}{\sum \alpha_i + 1}$$



## Simple Example: Dirichlet-Multinomial-Model

Data:	$\mathbf{k} \sim M(\boldsymbol{\theta})$	$(\sum k_j = n)$
conjugate prior:	$\boldsymbol{\theta} \sim \text{Dir}(\boldsymbol{\alpha})$	$(\sum \theta_j = 1)$
posterior:	$\boldsymbol{\theta} \mid \mathbf{k} \sim \text{Dir}(\boldsymbol{\alpha} + \mathbf{k})$	

$$\mathbb{E}[\theta_j] = \frac{\alpha_j}{\sum \alpha_i} =: y_j^{(0)} \quad \mathbb{V}(\theta_j) = \frac{\mathbb{E}[\theta_j](1 - \mathbb{E}[\theta_j])}{\sum \alpha_i + 1} = \frac{y_j^{(0)}(1 - y_j^{(0)})}{n^{(0)} + 1}$$

Data :	$\mathbf{k} \sim M(\boldsymbol{\theta})$	
conjugate prior:	$\boldsymbol{\theta} \sim \text{Dir}(n^{(0)}, \mathbf{y}^{(0)})$	$n^{(0)} = \sum \alpha_i$
posterior:	$\boldsymbol{\theta} \mid \mathbf{k} \sim \text{Dir}(n^{(1)}, \mathbf{y}^{(1)})$	



# Simple Example: Dirichlet-Multinomial-Model

Data:	$\mathbf{k} \sim M(\boldsymbol{\theta})$	$(\sum k_j = n)$
conjugate prior:	$\boldsymbol{\theta} \sim \text{Dir}(\boldsymbol{\alpha})$	$(\sum \theta_j = 1)$
posterior:	$\boldsymbol{\theta} \mid \mathbf{k} \sim \text{Dir}(\boldsymbol{\alpha} + \mathbf{k})$	

$$\mathbb{E}[\theta_j] = \frac{\alpha_j}{\sum \alpha_i} =: y_j^{(0)} \quad \mathbb{V}(\theta_j) = \frac{\mathbb{E}[\theta_j](1 - \mathbb{E}[\theta_j])}{\sum \alpha_i + 1} = \frac{y_j^{(0)}(1 - y_j^{(0)})}{n^{(0)} + 1}$$

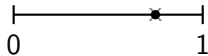
Data :	$\mathbf{k} \sim M(\boldsymbol{\theta})$	
conjugate prior:	$\boldsymbol{\theta} \sim \text{Dir}(n^{(0)}, \mathbf{y}^{(0)})$	$n^{(0)} = \sum \alpha_i$
posterior:	$\boldsymbol{\theta} \mid \mathbf{k} \sim \text{Dir}(n^{(1)}, \mathbf{y}^{(1)})$	

$$y_j^{(1)} = \frac{n^{(0)}}{n^{(0)} + n} \cdot y_j^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{k_j}{n} \quad n^{(1)} = n^{(0)} + n$$

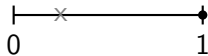


## Simple Example: Dirichlet-Multinomial-Model

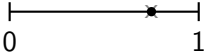
Case (i):  $y_j^{(0)} = 0.75,$   $k_j/n = 0.75$   
 $(n^{(0)} = 8)$   $(n = 16)$



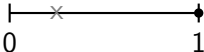
Case (ii):  $y_j^{(0)} = 0.25,$   $k_j/n = 1$   
 $(n^{(0)} = 8)$   $(n = 16)$



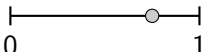
## Simple Example: Dirichlet-Multinomial-Model

Case (i):  $y_j^{(0)} = 0.75$ ,  $k_j/n = 0.75$  

$(n^{(0)} = 8)$   $(n = 16)$

Case (ii):  $y_j^{(0)} = 0.25$ ,  $k_j/n = 1$  

$(n^{(0)} = 8)$   $(n = 16)$

→  $\mathbb{E}[\theta_j | \mathbf{k}] = y_j^{(1)} = 0.75$ ,  $\mathbb{V}(\theta_j | \mathbf{k}) = 3/400$  

$(\mathbb{V}(\theta_j) = 1/48)$



Posterior inferences do not reflect uncertainty  
due to unexpected observations!





## Conjugate Priors

Weighted average structure is underneath *all common* conjugate priors for exponential family sampling distributions!

$X \stackrel{iid}{\sim}$  linear, canonical exponential family, i.e.

$$p(x | \theta) \propto \exp \{ \langle \psi, \tau(x) \rangle - n\mathbf{b}(\psi) \} \quad \left[ \psi \text{ transformation of } \theta \right]$$





## Conjugate Priors

Weighted average structure is underneath *all common* conjugate priors for exponential family sampling distributions!

$X \stackrel{iid}{\sim}$  linear, canonical exponential family, i.e.

$$p(x | \theta) \propto \exp \{ \langle \psi, \tau(x) \rangle - n\mathbf{b}(\psi) \} \quad \left[ \psi \text{ transformation of } \theta \right]$$

→ conjugate prior:

$$p(\theta) \propto \exp \{ n^{(0)} [\langle \psi, y^{(0)} \rangle - \mathbf{b}(\psi)] \}$$



## Conjugate Priors

Weighted average structure is underneath *all common* conjugate priors for exponential family sampling distributions!

$X \stackrel{iid}{\sim}$  linear, canonical exponential family, i.e.

$$p(x | \theta) \propto \exp \{ \langle \psi, \tau(x) \rangle - n\mathbf{b}(\psi) \} \quad \left[ \psi \text{ transformation of } \theta \right]$$

→ conjugate prior:

$$p(\theta) \propto \exp \{ n^{(0)} [\langle \psi, \mathbf{y}^{(0)} \rangle - \mathbf{b}(\psi)] \}$$

→ (conjugate) posterior:

$$p(\theta | x) \propto \exp \{ n^{(1)} [\langle \psi, \mathbf{y}^{(1)} \rangle - \mathbf{b}(\psi)] \},$$

where  $\mathbf{y}^{(1)} = \frac{n^{(0)}}{n^{(0)} + n} \cdot \mathbf{y}^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{1}{n} \tau(x)$  and  $n^{(1)} = n^{(0)} + n$ .



## Generalized iLUCK-models

Model for Bayesian inference with sets of priors  
(Walter & Augustin, 2009)

1. use conjugate priors from general construction method  
(prior parameters  $y^{(0)}$ ,  $n^{(0)}$ )



## Generalized iLUCK-models

Model for Bayesian inference with sets of priors  
(Walter & Augustin, 2009)

1. use conjugate priors from general construction method  
(prior parameters  $y^{(0)}$ ,  $n^{(0)}$ )
2. construct sets of priors via sets of parameters  
 $y^{(0)} \in \mathcal{Y}^{(0)} \times n^{(0)} \in \mathcal{N}^{(0)}$



## Generalized iLUCK-models

Model for Bayesian inference with sets of priors  
(Walter & Augustin, 2009)

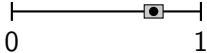
1. use conjugate priors from general construction method  
(prior parameters  $y^{(0)}$ ,  $n^{(0)}$ )
2. construct sets of priors via sets of parameters  
 $y^{(0)} \in \mathcal{Y}^{(0)} \times n^{(0)} \in \mathcal{N}^{(0)}$
3. set of posteriors  $\hat{=}$  set of (element-wise) updated priors  
 ➡ still easy to handle: described as set of  $(y^{(1)}, n^{(1)})$ 's

$$y^{(1)} = \frac{n^{(0)}}{n^{(0)} + n} \cdot y^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{1}{n} \tau(x)$$

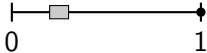
$$n^{(1)} = n^{(0)} + n$$

## Generalized iLUCK-models: Simple Example

Case (i):  $y_j^{(0)} \in [0.7, 0.8], \quad k_j/n = 0.75$   
 $(n^{(0)} \in [1, 8]) \quad (n = 16)$

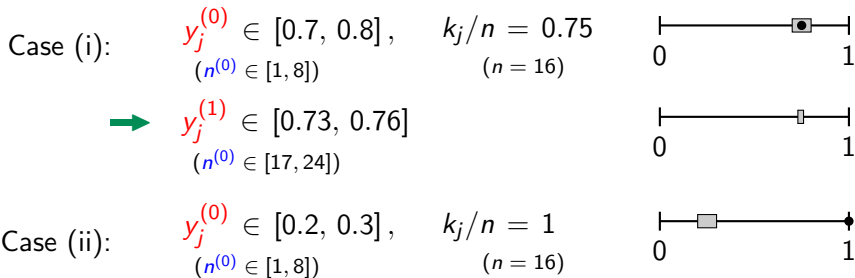


Case (ii):  $y_j^{(0)} \in [0.2, 0.3], \quad k_j/n = 1$   
 $(n^{(0)} \in [1, 8]) \quad (n = 16)$





## Generalized iLUCK-models: Simple Example



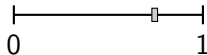


## Generalized iLUCK-models: Simple Example

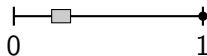
Case (i):  $y_j^{(0)} \in [0.7, 0.8]$ ,  $k_j/n = 0.75$   
 $(n^{(0)} \in [1, 8])$  ( $n = 16$ )



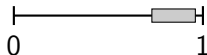
→  $y_j^{(1)} \in [0.73, 0.76]$   
 $(n^{(0)} \in [17, 24])$



Case (ii):  $y_j^{(0)} \in [0.2, 0.3]$ ,  $k_j/n = 1$   
 $(n^{(0)} \in [1, 8])$  ( $n = 16$ )



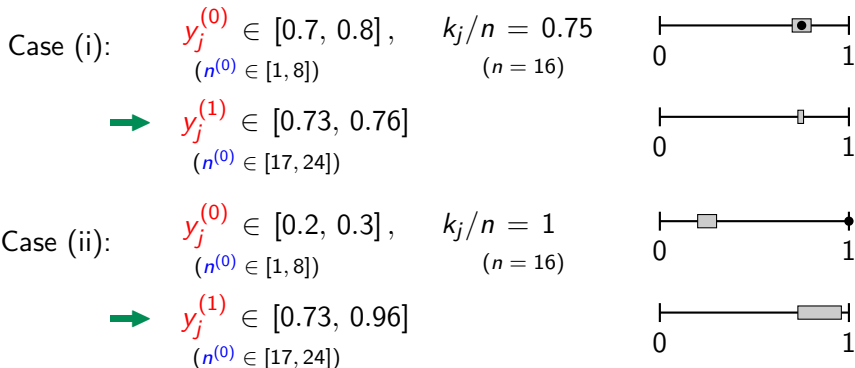
→  $y_j^{(1)} \in [0.73, 0.96]$   
 $(n^{(0)} \in [17, 24])$







## Generalized iLUCK-models: Simple Example



Generalized iLUCK-models lead to cautious inferences  
if, and only if, caution is needed.



## Generalized iLUCK-models: Parameter Sets

Prior parameter set  $y^{(0)} \in \mathcal{Y}^{(0)} \times n^{(0)} \in \mathcal{N}^{(0)}$  is rectangular,  
but posterior parameter set

$$\left\{ (y^{(1)}, n^{(1)}) \mid y^{(1)} = \frac{n^{(0)}y^{(0)} + \tau(x)}{n^{(0)} + n}, n^{(1)} = n^{(0)} + n, \right. \\ \left. y^{(0)} \in \mathcal{Y}^{(0)}, n^{(0)} \in \mathcal{N}^{(0)} \right\}$$

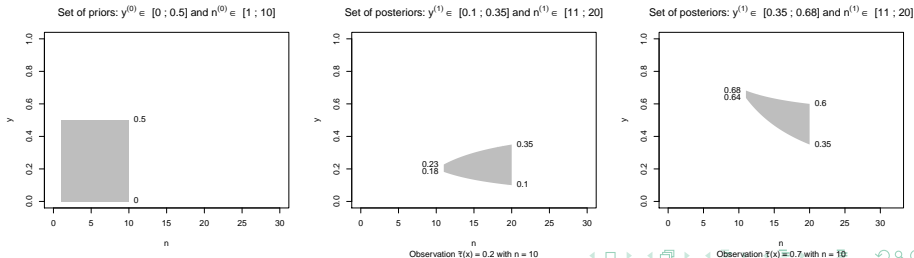
is not rectangular anymore!

## Generalized iLUCK-models: Parameter Sets

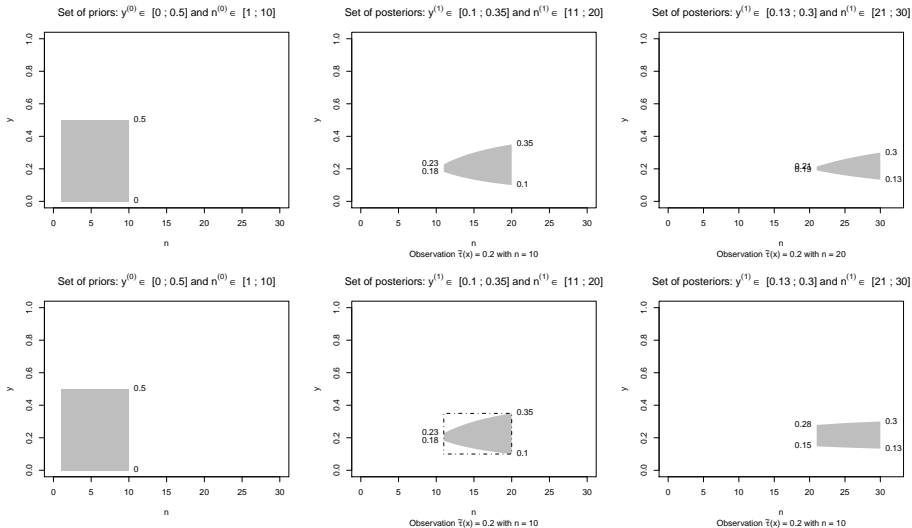
Prior parameter set  $y^{(0)} \in \mathcal{Y}^{(0)} \times n^{(0)} \in \mathcal{N}^{(0)}$  is rectangular,  
but posterior parameter set

$$\left\{ (y^{(1)}, n^{(1)}) \mid y^{(1)} = \frac{n^{(0)} y^{(0)} + \tau(x)}{n^{(0)} + n}, n^{(1)} = n^{(0)} + n, \right. \\ \left. y^{(0)} \in \mathcal{Y}^{(0)}, n^{(0)} \in \mathcal{N}^{(0)} \right\}$$

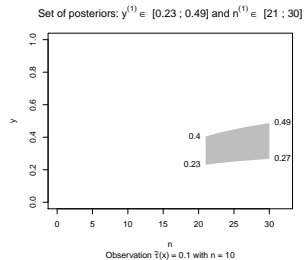
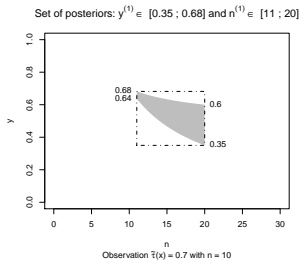
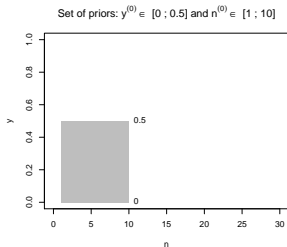
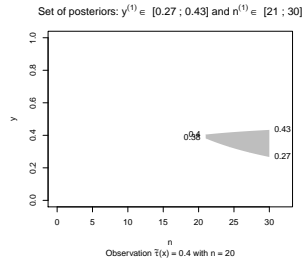
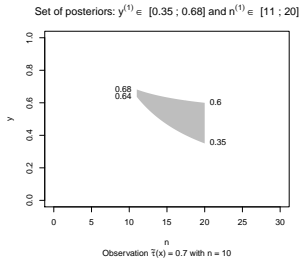
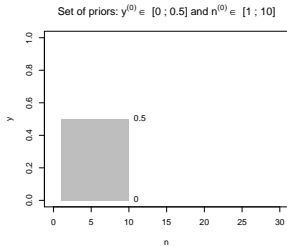
is not rectangular anymore!



# Generalized iLUCK-models: Direct vs. Sequential



# Generalized iLUCK-models: Direct vs. Sequential





# Generalized iLUCK-models: Direct vs. Sequential

with "rectangularization" . . .

- ▶ different intervals for  $y$  in case of prior-data conflict



## Generalized iLUCK-models: Direct vs. Sequential

with "rectangularization" . . .

- ▶ different intervals for  $y$  in case of prior-data conflict
- ▶ sufficient statistic  $\tau(x)$  not sufficient anymore?!?



## Generalized iLUCK-models: Direct vs. Sequential

with "rectangularization" . . .

- ▶ different intervals for  $y$  in case of prior-data conflict
- ▶ sufficient statistic  $\tau(x)$  not sufficient anymore?!?
- ▶ more information (order) → more imprecise posterior!





## Generalized iLUCK-models: Direct vs. Sequential

with "rectangularization" ...

- ▶ different intervals for  $y$  in case of prior-data conflict
- ▶ sufficient statistic  $\tau(x)$  not sufficient anymore?!?
- ▶ more information (order)  $\rightarrow$  more imprecise posterior!

In general:  $\mathcal{Y}_{\text{dir}}^{(2)} \subseteq \mathcal{Y}_{\text{seq}}^{(2)}$

if no pdc in both steps:

$$\mathcal{Y}_{\text{dir}}^{(2)} = \mathcal{Y}_{\text{seq}}^{(2)}$$

if pdc in both steps in same direction:

$$\mathcal{Y}_{\text{dir}}^{(2)} = \mathcal{Y}_{\text{seq}}^{(2)}$$

if neither of the former:

$$\mathcal{Y}_{\text{dir}}^{(2)} \subset \mathcal{Y}_{\text{seq}}^{(2)}$$

especially: if prior-data conflict 'on the way', then  $\mathcal{Y}_{\text{dir}}^{(2)} \subset \mathcal{Y}_{\text{seq}}^{(2)}$ !



## Generalized iLUCK-models: Direct vs. Sequential

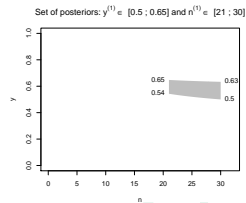
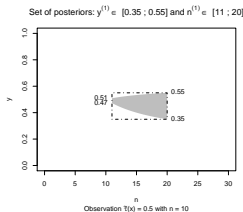
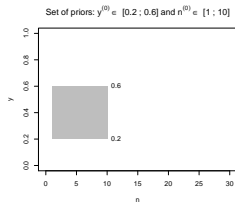
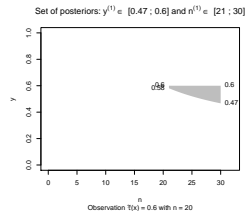
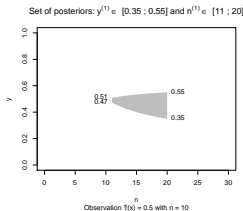
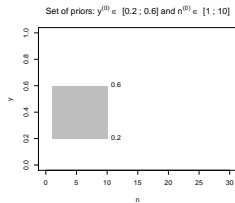
If  $\mathcal{Y}_{\text{dir}}^{(2)} \neq \mathcal{Y}_{\text{seq}}^{(2)}$ , there was prior-data conflict 'on the way'.

(but: no need to use  $\mathcal{Y}_{\text{seq}}^{(2)}$  for inference if you think, e.g., that the sequence of data is exchangeable nevertheless.)

# Generalized iLUCK-models: Direct vs. Sequential

If  $\mathcal{Y}_{\text{dir}}^{(2)} \neq \mathcal{Y}_{\text{seq}}^{(2)}$ , there was prior-data conflict 'on the way'.

(but: no need to use  $\mathcal{Y}_{\text{seq}}^{(2)}$  for inference if you think, e.g., that the sequence of data is exchangeable nevertheless.)





## Generalized iLUCK-models: Interval Lengths for 3 Steps

Data situation: Bernoulli sampling (observe 0 or 1)

Start with  $\mathcal{Y}^{(0)} = [0, 1]$

Compare sequences:

A (1, 0, 1)

B (1, 1, 0)

C (0, 1, 1)

D ( $[2/3, n = 3]$ ) (direct)



## Generalized iLUCK-models: Interval Lengths for 3 Steps

Data situation: Bernoulli sampling (observe 0 or 1)

Start with  $\mathcal{Y}^{(0)} = [0, 1]$

Compare sequences:

A (1, 0, 1)

B (1, 1, 0)

C (0, 1, 1)

D ( $[2/3, n = 3]$ ) (direct)

$A^* = C^*$  ( $[1/2, n = 2], 1$ )

$B^*$  ( $[1, n = 2], 0$ )



## Generalized iLUCK-models: Interval Lengths for 3 Steps

Data situation: Bernoulli sampling (observe 0 or 1)

Start with  $\mathcal{Y}^{(0)} = [0, 1]$

Compare sequences:

A (1, 0, 1)

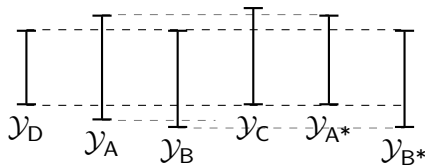
B (1, 1, 0)

C (0, 1, 1)

D ( $[2/3, n = 3]$ ) (direct)

$A^* = C^*$  ( $[1/2, n = 2], 1$ )

$B^*$  ( $[1, n = 2], 0$ )



Interval length for  $y$ : " $|\mathcal{Y}|$ " =:  $l$

$$l_A > l_B, l_A > l_C$$



## Generalized iLUCK-models: Interval Lengths for 3 Steps

Data situation: Bernoulli sampling (observe 0 or 1)

Start with  $\mathcal{Y}^{(0)} = [0, 1]$

Compare sequences:

A (1, 0, 1)

B (1, 1, 0)

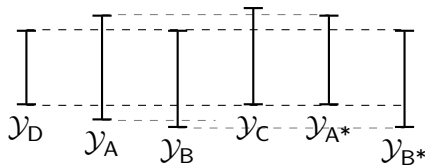
C (0, 1, 1)

D ( $[2/3, n = 3]$ ) (direct)

$A^* = C^*$  ( $[1/2, n = 2], 1$ )

$B^*$  ( $[1, n = 2], 0$ )     $A^{**} = B^{**}$  ( $1, [1/2, n = 2]$ )  $\hat{=}$  D (if  $\bar{n}^{(0)} \geq 1$ )

$C^{**}$  ( $0, [1, n = 2]$ )  $\hat{=}$  C



Interval length for  $y$ : " $|\mathcal{Y}|$ " =:  $l$

$l_A > l_B, l_A > l_C$

## “Strong Happiness”

Data situation: Bernoulli sampling (observe 0 or 1).

Idea: Choose sample size  $n_1$  such that  $I^{(1)}$  low enough under certain threshold  $I$  such that – even if we got another sample  $n_2$  in conflict to  $\mathcal{Y}^{(1)}$  –  $I^{(2)}$  is still below  $I$ !

With a sample of such a size  $n_1$  we attain “strong happiness”, because whatever we would see as a following sample, we would never get more imprecise than  $I$ !

(Is only possible because degree of prior-data conflict is bounded due to Bernoulli sampling!).





## "Strong Happiness"

$$I^{(2)} = \frac{\bar{n}^{(0)} I^{(0)}}{\bar{n}^{(0)} + n_1 + n_2} + \frac{\bar{n}^{(0)} - \underline{n}^{(0)}}{\bar{n}^{(0)} + n_1 + n_2} \left( \frac{n_1}{\underline{n}^{(0)} + n_1} \Delta\left(\frac{c_1}{n_1}, \mathcal{Y}^{(0)}\right) + \frac{n_2}{\underline{n}^{(0)} + n_1 + n_2} \Delta\left(\frac{c_2}{n_2}, \mathcal{Y}^{(1)}\right) \right)$$

$$\stackrel{!}{\leq} I \quad \forall (k_2, n_2)$$

## "Strong Happiness"

$$I^{(2)} = \frac{\bar{n}^{(0)} I^{(0)}}{\bar{n}^{(0)} + n_1 + n_2} + \frac{\bar{n}^{(0)} - \underline{n}^{(0)}}{\bar{n}^{(0)} + n_1 + n_2} \left( \frac{n_1}{\underline{n}^{(0)} + n_1} \Delta \left( \frac{c_1}{n_1}, \mathcal{Y}^{(0)} \right) + \frac{n_2}{\underline{n}^{(0)} + n_1 + n_2} \Delta \left( \frac{c_2}{n_2}, \mathcal{Y}^{(1)} \right) \right)$$

$$\stackrel{!}{\leq} I \quad \forall (k_2, n_2)$$

If data from second sample in conflict with  $\mathcal{Y}^{(1)}$  (and conflict strong enough), then imprecision  $I^{(2)}$  should first increase and then decrease in  $n_2$ .

→ find  $n_2$  that maximizes  $I^{(2)}$  for maximal possible conflict with current  $\mathcal{Y}^{(1)}$ , plug into formula for  $I^{(2)}$  and give  $n_1$  as a function of  $I^{(0)}$  (or  $\mathcal{Y}^{(0)}$ ) and  $I$  (and  $\mathcal{N}^{(0)}$ ).



## Plans, Further Ideas

- ▶ Try to attain Strong Happiness by sorting out formula
- ▶ Write programs for Strong Happiness
- ▶ Investigate interval lengths in general situation (not only Bernoulli sampling)
- ▶ Focus considerations on certain event of interest instead of  $\mathcal{Y}$ ?
- ▶ ...